

Rěšené pŕíklady na maxima a minima

Definice. Necht' Y_1, \dots, Y_n, Y jsou náhodné veličiny. Pak $Y_n \xrightarrow{P} Y$, pokud $\forall \varepsilon > 0: P[|Y_n - Y| > \varepsilon] \xrightarrow{n \rightarrow \infty} 0$

Tvrzení 1. $Y_n \xrightarrow{P} Y \Rightarrow Y_n \xrightarrow{d} Y; Y_n \xrightarrow{d} c \Rightarrow Y_n \xrightarrow{P} c$

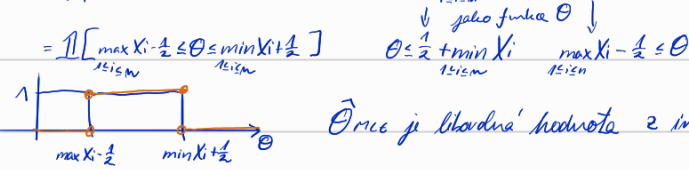
Tvrzení 2. Necht' X_1, \dots, X_n je náhodný výběr z rozdělení s CDF $F_X(x), x \in \mathbb{R}$. Pak CDF n. veličiny $Y_n = \max_{1 \leq i \leq n} X_i$ je $F_{Y_n}(y) = F_X^n(y)$ a CDF min. veličiny

$Z_n = \min_{1 \leq i \leq n} X_i$ je $F_{Z_n}(z) = 1 - (1 - F_X(z))^n$

Všechny systémy hustot jsou neregulární (mohou záviset na parametru)

Příklad 55 $X_1, \dots, X_n \sim \text{Unif}(\theta - \frac{1}{2}, \theta + \frac{1}{2})$ s hustotou $f(x, \theta) = \mathbb{1}[\theta - \frac{1}{2} \leq x \leq \theta + \frac{1}{2}]$. MLE + slabá konzistence.

Rěšení $L_n(\theta) = \prod_{i=1}^n \mathbb{1}[\theta - \frac{1}{2} \leq X_i \leq \theta + \frac{1}{2}] = \mathbb{1}[\theta - \frac{1}{2} \leq \min_{1 \leq i \leq n} X_i \leq \max_{1 \leq i \leq n} X_i \leq \theta + \frac{1}{2}]$



$\hat{\theta}_{MLE}$ je libovolná hodnota z intervalu $[\max X_i - \frac{1}{2}, \min X_i + \frac{1}{2}]$

interval je nejvíce rozptýlený, protože $0 \leq \max_{1 \leq i \leq n} X_i - \min_{1 \leq i \leq n} X_i \leq 1$

Chceme ukázat $\min X_i \xrightarrow{P} \theta - \frac{1}{2}, \max X_i \xrightarrow{P} \theta + \frac{1}{2}$, pak lib. hodnota z $[\max X_i - \frac{1}{2}, \min X_i + \frac{1}{2}]$ konv. v P k θ , protože krajní hodnoty $\xrightarrow{P} \theta$.

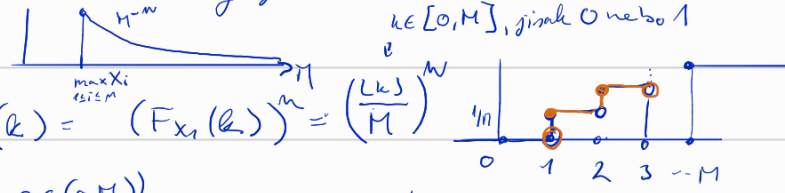
$$F_{X_i}(x) = \begin{cases} 0 & x < \theta - \frac{1}{2} \\ x - \theta + \frac{1}{2} & \theta - \frac{1}{2} \leq x \leq \theta + \frac{1}{2} \\ 1 & x > \theta + \frac{1}{2} \end{cases} \rightsquigarrow \text{Tvrzení 2.} \rightsquigarrow F_{\min}(x) = \begin{cases} 0 & x < \theta - \frac{1}{2} \\ 1 - (\theta + \frac{1}{2} - x)^n & \theta - \frac{1}{2} \leq x \leq \theta + \frac{1}{2} \\ 1 & x > \theta + \frac{1}{2} \end{cases}$$

ukážeme pak $\lim_{n \rightarrow \infty} F_{\min}(x) = \begin{cases} 0 & x < \theta - \frac{1}{2} \\ 1 & x > \theta - \frac{1}{2} \end{cases}$, což je po upravení bodu nepojistit CDF degenerované n.v. $\theta - \frac{1}{2}$

$(\min \xrightarrow{d} \theta - \frac{1}{2} \Rightarrow \min \xrightarrow{P} \theta - \frac{1}{2})$ analogicky $\lim_{n \rightarrow \infty} F_{\max}(x) = \begin{cases} 0 & x < \theta + \frac{1}{2} \\ 1 & x > \theta + \frac{1}{2} \end{cases}$

Příklad 56. $X_1, \dots, X_n \sim \text{Unif}\{1, \dots, M\}$ $f(x, M) = \frac{1}{M} \mathbb{1}\{x \in \{1, \dots, M\}\}$ (mereg. sgd) MLE + konzistence M

Rěšení. $L_n(M) = M^{-n} \cdot \mathbb{1}[\max_{1 \leq i \leq n} X_i \leq M]$



$\hat{M}_{MLE} = \max_{1 \leq i \leq n} X_i \rightsquigarrow$ rozdělení $F_{\hat{M}}(k) = (F_{X_1}(k))^n = (\frac{k}{M})^n$

slabá konzistence: zvolme $\varepsilon > 0$ (stačí $\varepsilon \in (0, M)$)

$$P[|M - \hat{M}| > \varepsilon] = P[\hat{M} < M - \varepsilon] \leq P[\hat{M} \leq M - \varepsilon] = (\frac{M - \varepsilon}{M})^n = (1 - \frac{\varepsilon}{M})^n \rightarrow 0 \quad \forall \varepsilon > 0$$

$$F(x; \lambda, \delta) = \begin{cases} 0 & x < \delta \\ 1 - e^{-\lambda(x-\delta)} & x \geq \delta \end{cases}$$

Příklad 67 X_1, \dots, X_n iid $f(x; \lambda, \delta) = \lambda e^{-\lambda(x-\delta)} \mathbb{1}[x \in (\delta, \infty)]$, MLE, slabá konzistence, $P[M(\hat{\delta} - \delta) \leq x]$

Rěšení MLE $L_n(\lambda, \delta) = \lambda^n \exp\{-\lambda(\sum X_i - n\delta)\} \mathbb{1}[\delta < \min_{1 \leq i \leq n} X_i]$

pro první 1 $\rightsquigarrow \hat{\delta}_{MLE} = \min X_i$ a pro druhé δ je $\hat{\lambda}_{MLE}$ řešení $\frac{\partial}{\partial \lambda} -\sum X_i + n\delta = 0 \rightsquigarrow \hat{\lambda} = \frac{1}{\bar{X}_n - \hat{\delta}}$

$(\hat{\delta}_n, \hat{\lambda}_n) = (\min X_i, \frac{1}{\bar{X}_n - \hat{\delta}}) \xrightarrow{P} (\delta, \lambda)$

$F_{\hat{\delta}}(y) \stackrel{T_2}{=} \begin{cases} 0 & y < \delta \\ 1 - (e^{-\lambda(y-\delta)})^n & y \geq \delta \end{cases} \Rightarrow \hat{\delta} - \delta \sim \text{Exp}(n\lambda) \text{ a } n(\hat{\delta} - \delta) \sim \text{Exp}(1)$

$$P[n(\hat{\sigma} - \sigma) \leq x] = P[\hat{\sigma} - \sigma \leq \frac{x}{n}] \underset{\uparrow}{=} 1 - e^{-\ln(\frac{x}{n})} = 1 - e^{-1x}$$

mijen asymptoticky

Příklad 68 $X_1, \dots, X_n \sim \text{Unif}(a, b)$ MLE, sl. konzistence, $P[n(\hat{b} - b) \leq x]$

Řešení! $L_n(a, b) = \left(\frac{1}{b-a}\right)^n \mathbb{1}[a < \min X_i < \max X_i < b]$

$(\hat{a}, \hat{b}) = (\min X_i, \max X_i) \rightarrow$ stabilní konzistence

$F_X(x) = \begin{cases} \frac{x-a}{b-a} & x \in (a, b) \\ 0 & x \leq a \\ 1 & x \geq b \end{cases} \rightsquigarrow F_{\hat{b}}(x) = \begin{cases} \left(\frac{x-a}{b-a}\right)^n & x \in (a, b) \\ 0 & x \leq a \\ 1 & x \geq b \end{cases} \lim_{n \rightarrow \infty} F_{\hat{b}}(x) = \begin{cases} 1 & x \geq b \\ 0 & x < b \end{cases} \rightarrow \hat{b} \xrightarrow{d} b \Rightarrow \hat{b} \xrightarrow{P} b$

konvergence i slabě, v P

$F_{\hat{a}}(x) = \begin{cases} 1 - \left(\frac{b-x}{b-a}\right)^n & x \in (a, b) \\ 0 & x \leq a \\ 1 & x \geq b \end{cases} \lim_{n \rightarrow \infty} F_{\hat{a}}(x) = \begin{cases} 1 & x > a \\ 0 & x \leq a \end{cases} + \text{mod. v bodě nespojivosti} \hat{a} \xrightarrow{d} a \Rightarrow \hat{a} \xrightarrow{P} a \quad \left(\frac{\hat{a}}{\hat{b}}\right)^P \rightarrow \left(\frac{a}{b}\right)$

$P[n(\hat{b} - b) \leq x] = P[\hat{b} \leq \frac{x}{n} + b] = \left(\frac{\frac{x}{n} + b - a}{b - a}\right)^n = \left(1 + \frac{x}{n(b-a)}\right)^n \xrightarrow{n \rightarrow \infty}$

$\rightarrow n(b - \hat{b}) \xrightarrow{d} \text{Exp}\left(\frac{1}{b-a}\right) \rightarrow e^{\frac{x}{b-a}}$ pro $x < 0$