

Breusch-Pagan's test - complete version

$(\begin{smallmatrix} x_1 \\ y_1 \end{smallmatrix}), \dots, (\begin{smallmatrix} x_m \\ y_m \end{smallmatrix}) \sim \text{iid normal linear model, } \text{var}(y_i | x_i) = \sigma^2 \tau_i(\beta, \lambda)$

τ_i given functions $\mathbb{R} \rightarrow [0, \infty)$ that do not dep. on y $E(y_i | x_i) = x_i' \beta$ $\theta = (\beta^T, \sigma^2, \lambda^T) \in \mathbb{R}^{k+1+k}$
 $\tau_i(\beta, 0) = 1$ Rao's score test $H_0: \lambda = 0, H_1: \lambda \neq 0$.

$$L(\theta) = c \frac{1}{\sigma^m \prod \tau_i} \exp \left\{ -\frac{1}{2\sigma^2} \sum \frac{(y_i - x_i' \beta)^2}{\tau_i} \right\} \quad e_i := y_i - x_i' \beta$$

$$l(\theta) = \log c - \frac{m}{2} \log \sigma^2 - \frac{1}{2} \sum \log \tau_i - \frac{1}{2\sigma^2} \sum \frac{e_i^2}{\tau_i}$$

$$\frac{\partial l(\theta)}{\partial \beta} = -\frac{1}{2} \sum \frac{\partial \log \tau_i}{\partial \beta} + \frac{1}{2\sigma^2} \sum \frac{2e_i x_i}{\tau_i} + \frac{1}{2\sigma^2} \sum \left(\frac{e_i}{\tau_i} \right)^2 \frac{\partial \log \tau_i}{\partial \beta}$$

$$= -\frac{1}{\sigma^2} \sum \frac{e_i x_i}{\tau_i} + \frac{1}{2} \sum \left(\left(\frac{e_i}{\sigma \tau_i} \right)^2 - 1 \right) \frac{\partial \log \tau_i}{\partial \beta}$$

$$\frac{\partial}{\partial \beta} \left(\frac{e_i}{\sigma \tau_i} \right)^2 = \frac{2}{\sigma} \left(\frac{e_i}{\sigma \tau_i} \right) \cdot \left(-x_i' \tau_i^{-1} - \frac{e_i}{2\tau_i} \frac{\partial \tau_i}{\partial \beta} \right)$$

$$= - \left(\frac{2e_i x_i}{\tau_i \sigma^2} + \frac{e_i^2}{\tau_i^2 \sigma^2} \frac{\partial \tau_i}{\partial \beta} \right)$$

$$\frac{\partial l(\theta)}{\partial \sigma^2} = -\frac{m}{2\sigma^2} + \frac{1}{2\sigma^4} \sum \frac{e_i^2}{\tau_i} = -\frac{m}{2\sigma^2} + \frac{1}{2\sigma^2} \sum \left(\frac{e_i}{\sigma \tau_i} \right)^2$$

$$\frac{\partial}{\partial \lambda} \left(\frac{e_i}{\sigma \tau_i} \right)^2 = -\frac{e_i^2}{\sigma^2} \cdot \frac{1}{\tau_i^2} \frac{\partial \tau_i}{\partial \lambda} = - \left(\frac{e_i}{\sigma \tau_i} \right)^2 \frac{\partial \log \tau_i}{\partial \lambda}$$

$$\frac{\partial l(\theta)}{\partial \lambda} = -\frac{1}{2} \sum \frac{\partial \log \tau_i}{\partial \lambda} + \frac{1}{2} \sum \left(\frac{e_i}{\sigma \tau_i} \right)^2 \frac{\partial \log \tau_i}{\partial \lambda}$$

$$= \frac{1}{2} \sum \left(\left(\frac{e_i}{\sigma \tau_i} \right)^2 - 1 \right) \frac{\partial \log \tau_i}{\partial \lambda}$$

$$\frac{\partial}{\partial \beta^T} \frac{e_i}{\tau_i} = \frac{-x_i' \tau_i - e_i \frac{\partial \tau_i}{\partial \beta^T}}{\tau_i^2} = - \left(\frac{x_i}{\tau_i} + \frac{e_i}{\tau_i^2} \frac{\partial \tau_i}{\partial \beta^T} \right)$$

$$\frac{\partial}{\partial \lambda^T} \frac{e_i}{\tau_i} = -\frac{e_i}{\tau_i^2} \frac{\partial \tau_i}{\partial \lambda^T}$$

$$\frac{\partial^2 l(\theta)}{\partial \beta \partial \beta^T} = -\frac{1}{\sigma^2} \sum \frac{x_i x_i'}{\tau_i} - \frac{1}{\sigma^2} \sum \frac{x_i e_i}{\tau_i} \frac{\partial \log \tau_i}{\partial \beta^T} + \frac{1}{2} \sum \left(-\frac{2e_i x_i'}{\tau_i \sigma^2} - \frac{e_i^2}{\sigma^2 \tau_i} \frac{\partial \log \tau_i}{\partial \beta^T} \right)$$

vanishes in expectation
 $+ \frac{1}{2} \sum \left(\left(\frac{e_i}{\sigma \tau_i} \right)^2 - 1 \right) \frac{\partial^2 \log \tau_i}{\partial \beta \partial \beta^T}$

$$\frac{\partial^2 l(\theta)}{\partial \beta \partial \sigma^2} = -\frac{1}{2\sigma^4} \sum \frac{e_i x_i'}{\tau_i} - \frac{1}{2\sigma^4} \sum \frac{e_i^2}{\tau_i} \frac{\partial \log \tau_i}{\partial \beta}$$

$$\frac{\partial^2 l(\theta)}{\partial \beta \partial \lambda^T} = -\frac{1}{\sigma^2} \sum \frac{x_i e_i}{\tau_i} \frac{\partial \log \tau_i}{\partial \lambda^T} + \frac{1}{2} \sum \frac{\partial \log \tau_i}{\partial \beta} \cdot \left(-\left(\frac{e_i}{\sigma \tau_i} \right)^2 \frac{\partial \log \tau_i}{\partial \lambda^T} \right) + \frac{1}{2} \sum \left(\left(\frac{e_i}{\sigma \tau_i} \right)^2 - 1 \right) \frac{\partial^2 \log \tau_i}{\partial \beta \partial \lambda^T}$$

$$\frac{\partial^2 l(\theta)}{\partial \sigma^2 \partial \sigma^2} = \frac{m}{2\sigma^4} - \frac{1.2}{2\sigma^6} \sum \left(\frac{e_i}{\tau_i} \right)^2$$

$$\frac{\partial^2 l(\theta)}{\partial \sigma^2 \partial \lambda^T} = -\frac{1}{2\sigma^2} \sum \left(\frac{e_i}{\sigma \tau_i} \right)^2 \frac{\partial \log \tau_i}{\partial \lambda^T}$$

$$\frac{\partial^2 l(\theta)}{\partial \lambda \partial \lambda^T} = \frac{1}{2} \sum \frac{\partial \log \tau_i}{\partial \lambda} \cdot \left(-\left(\frac{e_i}{\sigma \tau_i} \right)^2 \frac{\partial \log \tau_i}{\partial \lambda^T} \right) + \frac{1}{2} \sum \left(\left(\frac{e_i}{\sigma \tau_i} \right)^2 - 1 \right) \frac{\partial^2 \log \tau_i}{\partial \lambda \partial \lambda^T}$$

$E \left(\frac{e_i}{\sigma \tau_i} \right)^2 = 1$ $E e_i = 0 \Rightarrow$ Fisher information matrix

$$J_m(\theta) = \begin{pmatrix} \frac{1}{\sigma^2} X' W X + \frac{1}{2} D_\beta' D_\beta & \frac{1}{2\sigma^2} D_\beta' \mathbb{1} & \frac{1}{2} D_\beta' D_\lambda \\ \frac{1}{2\sigma^2} \mathbb{1}' D_\beta & m/2\sigma^4 & \frac{1}{2\sigma^2} \mathbb{1}' D_\lambda \\ \frac{1}{2} D_\lambda' D_\beta & \frac{1}{2\sigma^2} D_\lambda' \mathbb{1} & \frac{1}{2} D_\lambda' D_\lambda \end{pmatrix}$$

$W = \text{diag } \tau_i^{-1}$
 $D_\beta = \frac{\partial \log \tau}{\partial \beta^T} \leftarrow m \times k \text{ matrix}$
 $D_\lambda = \frac{\partial \log \tau}{\partial \lambda^T} \leftarrow m \times k \text{ matrix}$

under H_0 : $\tilde{\lambda} = 0$, $\tilde{\beta} = (X'X)^{-1}X'Y$, $\tilde{\sigma}^2 = \frac{1}{m} \sum \tilde{e}_i^2 = \frac{1}{2} \sum (Y_i - X_i \tilde{\beta})^2$ (slam norm homos)

$$\Rightarrow \frac{\partial \ell(\theta)}{\partial \theta} \Big|_{(\tilde{\beta}, \tilde{\sigma}^2, \tilde{\lambda})} = \left(0^T, 0, \frac{1}{2} \left[\left(\frac{\tilde{e}_i}{\tilde{\sigma}} \right)^2 - 1 \right] \frac{\partial \log \tau_i}{\partial \lambda} \Big|_{\lambda=0} \right)^T$$

$$R_m := \frac{\partial \ell(\theta)}{\partial \theta^T} \Big|_{\tilde{\beta}, \tilde{\sigma}^2, 0} \cdot [J_m(\tilde{\theta})]^{-1} \frac{\partial \ell(\theta)}{\partial \theta} \Big|_{\tilde{\beta}, \tilde{\sigma}^2, 0} \xrightarrow{H_0} \chi^2$$

• Special case: $\tau_i(\beta, \lambda) = \exp\{\lambda \cdot x_i' \beta\} = \exp\{\lambda \cdot E[Y_i | X_i]\}$ for $\lambda \in \mathbb{R}$

Then $D_\lambda = X\beta$, $D_\beta = \lambda X$, under H_0 : $\lambda=0$ we have $D_\beta = 0_{m \times 2}$, and

$$J_m = \left(\begin{array}{cc|c} \frac{1}{\sigma^2} X'X & 0 & 0 \\ 0' & \frac{m}{2\sigma^4} & \frac{11'X\beta}{2\sigma^2} \\ \hline 0' & \frac{\beta'X'11}{2\sigma^2} & \frac{1}{2} \beta'X'X\beta \end{array} \right) \quad \frac{\partial \ell(\theta)}{\partial \theta} \Big|_{\tilde{\theta}} = \left(0^T, 0, \frac{1}{2\tilde{\sigma}^2} \sum (\tilde{e}_i^2 - \tilde{\sigma}^2) \hat{Y}_i \right)^T$$

and we need only the $(k+2, k+2)$ -element of $J_m(\tilde{\theta})^{-1}$. Invert the partitioned matrix

$$\text{using } \begin{pmatrix} A & B \\ C & D \end{pmatrix}^{-1} = \begin{pmatrix} \cdot & \cdot \\ \cdot & (D - CA^{-1}B)^{-1} \end{pmatrix} \text{ to get}$$

$$[J_m(\tilde{\theta})]_{k+2, k+2}^{-1} = \left[\frac{1}{2} \tilde{\beta}' X' X \tilde{\beta} - (0' \frac{\tilde{\beta}' X' 11}{2\tilde{\sigma}^2}) \begin{pmatrix} \sigma^2 (X'X)^{-1} & 0 \\ 0' & \frac{2\tilde{\sigma}^4}{m} \end{pmatrix} \begin{pmatrix} 0 \\ \frac{11' X \tilde{\beta}}{2\tilde{\sigma}^2} \end{pmatrix} \right]^{-1}$$

$$= \left(\frac{1}{2} \hat{Y}' \hat{Y} - \frac{1}{2m} (\hat{Y}' 11)^2 \right)^{-1} = \left(\frac{1}{2} \sum (\hat{Y}_i - \bar{\hat{Y}})^2 \right)^{-1}$$

Because $\sum \tilde{e}_i^2 = m \tilde{\sigma}^2$ we can write $\sum (\tilde{e}_i^2 - \tilde{\sigma}^2) \hat{Y}_i = \sum \tilde{e}_i^2 \hat{Y}_i - \frac{1}{m} \sum \tilde{e}_i^2 \cdot \sum \hat{Y}_i$ and

$$= \sum \tilde{e}_i^2 (\hat{Y}_i - \bar{\hat{Y}})$$

$$R_m = \frac{1}{2(\tilde{\sigma}^2)^2} \cdot \frac{(\sum (Y_i - X_i' \tilde{\beta})^2 (X_i' \tilde{\beta} - \bar{\hat{Y}}))^2}{\sum (\hat{Y}_i - \bar{\hat{Y}})^2} = \frac{(\sum (Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{\hat{Y}}))^2}{2(\tilde{\sigma}^2)^2 \sum (\hat{Y}_i - \bar{\hat{Y}})^2} \xrightarrow{H_0} \chi^2$$

Example 19. $X_1, \dots, X_m \stackrel{iid}{\sim} \text{Mult}(1, p)$ $P = (p_1, \dots, p_k)^T$ $p_k = 1 - \sum_{z=1}^{k-1} p_z$

$H_0: P = P_0$ $H_1: P \neq P_0$

$L(P) = \prod_{i=1}^m \left(\prod_{z=1}^{k-1} p_z^{x_{iz}} \right) \cdot \left(1 - \sum_{z=1}^{k-1} p_z \right)^{1 - \sum_{z=1}^{k-1} x_{iz}}$ $N_k := \sum_{i=1}^m x_{ik}$

$\ell(P) = \sum N_k \log p_k + N_k \log (1 - \sum p_z)$

$\frac{\partial \ell(P)}{\partial P} = \left(\frac{N_k}{p_k} - \frac{N_k}{1 - \sum p_z} \right)_{z=1}^{k-1} \stackrel{!}{=} 0 \Rightarrow p_k = 1 - \sum_{z=1}^{k-1} p_z$

$\left[\frac{\partial^2 \ell(P)}{\partial p \partial p^T} \right]_{z, \ell} = \begin{cases} -\frac{N_k}{p_k^2} - \frac{N_k}{(p_k)^2} & z = \ell \\ -\frac{N_k}{(p_k)^2} & z \neq \ell \end{cases}$

$\frac{N_k}{p_k} - \frac{N_k}{p_k} = 0$

$p_k = p_k \frac{m N_k}{N_k} \quad \Big| \sum_{z=1}^{k-1}$

$1 - p_k = p_k \left(\frac{m - N_k}{N_k} \right)$
 $\hat{p}_k = \frac{N_k}{m} \Rightarrow \hat{p}_k = \frac{N_k}{N}$

$\frac{\partial^2 \ell(P)}{\partial p \partial p^T} = -\text{diag} \left(\frac{N_k}{p_k^2} \right) - \mathbb{1} \mathbb{1}^T \cdot \frac{N_k}{p_k^2}$

$I(P) = \left[\text{diag} \left(\frac{m}{p_k} \right) + \mathbb{1} \mathbb{1}^T \frac{m}{p_k} \right]^{-1}$

LR: $= 2(\ell(\hat{P}) - \ell(P_0)) = 2 \left(\sum N_k \log \hat{p}_k + N_k \log \left(\frac{N_k}{m} \right) - \sum N_k \log p_{0k} - N_k \log p_{0k} \right)$
 $= 2 \left(\sum_{z=1}^k N_k \log \frac{N_k}{m p_{0z}} \right) \stackrel{\text{as } N_k}{\sim} \chi_{k-1}^2$

$I(P_0) \stackrel{MM}{=} \text{diag} \left(\frac{m}{p_{0k}} \right) + \mathbb{1} \mathbb{1}^T \frac{m}{p_{0k}}$

Sherman-Morrison formula

$(A + \mu \nu^T)^{-1} = A^{-1} - \frac{A^{-1} \mu \nu^T A^{-1}}{1 + \nu^T A^{-1} \mu}$

$\frac{\partial \ell(P_0)}{\partial P} = \left(\frac{N_k}{p_{0k}} - \frac{N_k}{p_{0k}} \right)_{z=1}^{k-1}$

$A = \text{diag} \left(\frac{m}{p_{0k}} \right) \quad \mu = \nu = \mathbb{1} \cdot \sqrt{\frac{m}{p_{0k}}}$

$R_m := \frac{\partial \ell(P_0)}{\partial P}^T I(P_0)^{-1} \frac{\partial \ell(P_0)}{\partial P}$

$I(P_0)^{-1} = \text{diag} \left(\frac{p_{0k}}{m} \right) - \frac{\text{diag} \left(\frac{p_{0k}}{m} \right) \frac{m}{p_{0k}} \mathbb{1} \mathbb{1}^T \text{diag} \left(\frac{p_{0k}}{m} \right)}{1 + \frac{m}{p_{0k}} \mathbb{1}^T \text{diag} \left(\frac{p_{0k}}{m} \right) \mathbb{1}}$

$= \left(\frac{N_k - N_k}{p_{0k}} \right) \left(\text{diag} \left(\frac{p_{0k}}{m} \right) - \frac{p_{0k} p_{0k}^T}{m} \right) \left(\frac{N_k}{p_{0k}} - \frac{N_k}{p_{0k}} \right)$

$= \text{diag} \left(\frac{p_{0k}}{m} \right) - \frac{p_{0k}}{p_{0k}} \left(\frac{p_{0k} p_{0k}}{m^2} \right)_{z, \ell=1}^{k-1} \frac{1}{1 + \frac{m}{p_{0k}} \sum \frac{p_{0k}}{m}}$

$= \frac{1}{m} \left[\sum \left(\frac{N_k}{p_{0k}} - \frac{N_k}{p_{0k}} \right)^2 \frac{p_{0k}}{m} - \left(\sum \left(\frac{N_k}{p_{0k}} - \frac{N_k}{p_{0k}} \right) p_{0k} \right)^2 \right]$

$= \text{diag} \left(\frac{p_{0k}}{m} \right) - \left(\frac{p_{0k} p_{0k}}{m} \right)_{z, \ell=1}^{k-1}$

$N = \begin{pmatrix} N_1 \\ \vdots \\ N_{k-1} \end{pmatrix}$
 $= \frac{1}{m} (I - II)$

$II = \sum \left(\frac{N_k}{p_{0k}} - \frac{N_k}{p_{0k}} \right) p_{0k} = \sum N_k - \frac{N_k}{p_{0k}} \sum p_{0k} = m - N_k - \frac{N_k}{p_{0k}} (1 - p_{0k}) = m - N_k + N_k - \frac{N_k}{p_{0k}} =$

$= \frac{m p_{0k} - N_k}{p_{0k}}$

19 contd. $I = \sum \left(\frac{N_k}{P_{0k}} - \frac{N_k}{P_{0k}} - m + m \right)^2 P_{0k} = \sum \frac{(N_k - m P_{0k})^2}{P_{0k}} + \sum \frac{(N_k - m P_{0k})^2}{P_{0k}^2} P_{0k}$

$- \sum 2 \frac{N_k - m P_{0k}}{P_{0k}} \cdot \frac{N_k - m P_{0k}}{P_{0k}} \cdot P_{0k} = \sum \frac{(N_k - m P_{0k})^2}{P_{0k}} + \frac{(N_k - m P_{0k})^2}{P_{0k}^2} (1 - P_{0k}) -$

$- 2 \frac{N_k - m P_{0k}}{P_{0k}} (m - N_k - m(1 - P_{0k})) = \sum \frac{(N_k - m P_{0k})^2}{P_{0k}} + \frac{(N_k - m P_{0k})^2}{P_{0k}} \left(\frac{1}{P_{0k}} - 1 + 2 \right)$

R_m $= \frac{1}{m} (I - II^2) = \frac{1}{m} \left(\sum \frac{(N_k - m P_{0k})^2}{P_{0k}} + \frac{(N_k - m P_{0k})^2}{P_{0k}} \left(\frac{1}{P_{0k}} + 1 - \frac{1}{P_{0k}} \right) \right)$

$= \frac{1}{m} \sum_{k=1}^k \frac{(N_k - m P_{0k})^2}{P_{0k}} \underset{H_0}{\sim} \chi_{k-1}^2$

W_m $= \left(\frac{N}{m} - P_0 \right)' I \left(\frac{N}{m} \right) \left(\frac{N}{m} - P_0 \right) = \left(\frac{N}{m} - P_0 \right)' \left(\text{diag} \left(\frac{m^2}{N} \right) + 111^T \frac{m^2}{N_k} \right) \left(\frac{N}{m} - P_0 \right)$

$= m^2 \cdot \left[\sum \frac{(N_k - m P_{0k})^2}{N_k m^2} + \frac{1}{N_k} \left[\sum \left(\frac{N_k}{m} - P_{0k} \right) \right]^2 \right] = \sum \frac{(N_k - m P_{0k})^2}{N_k} + \frac{m^2}{N_k} \left(\frac{m - N_k}{m} - (1 - P_{0k}) \right)^2$

$= \sum \frac{(N_k - m P_{0k})^2}{N_k} + \frac{m^2}{N_k} \frac{(N_k - m P_{0k})^2}{m^2} = \sum_{k=1}^k \frac{(N_k - m P_{0k})^2}{N_k} \underset{H_0}{\sim} \chi_{k-1}^2$