

- Current age: $\lambda(t) = \lambda_0(t) \cdot \exp\{\alpha \cdot (\text{age} + t) + \beta \cdot Z\}$, where age is the age at the entry to the study
 $\lambda(t|Z) = \lambda_0(t) \cdot \exp\{\alpha \cdot (\text{age} + t) + \beta^T Z\} = \lambda_0(t) \cdot e^{\alpha t} \cdot \exp\{\alpha \cdot \text{age} + \beta^T Z\}$
 and now compare it with model where only age itself is in the model: $\tilde{\lambda}(t|Z) = \tilde{\lambda}_0(t) \cdot \exp\{\tilde{\alpha} \cdot \text{age} + \tilde{\beta}^T Z\}$
 $\text{age} = 0, Z = 0 \Rightarrow \lambda_0(t) \cdot e^{\alpha t} \text{ vs. } \tilde{\lambda}_0(t)$, no interpretation of baseline hazard changes
 ↳ interpretation of λ and α is tied together
- $$\frac{\tilde{\lambda}(t|Z)}{\lambda(t|Z)} \stackrel{Z=0}{=} \exp\{\tilde{\alpha} \cdot (\text{age}_1 - \text{age}_2)\} \rightsquigarrow \tilde{\alpha} = \tilde{\lambda}(t|\text{age}+1, Z) / \tilde{\lambda}(t|\text{age}, Z)$$

$$\frac{\tilde{\lambda}(t|Z)}{\lambda(t|Z)} \stackrel{Z=0}{=} \exp\{\alpha \cdot (\text{age}_1 - \text{age}_2)\} \rightsquigarrow \alpha = \lambda(t|\text{age}+1, Z) / \lambda(t|\text{age}, Z)$$

The same interpretation
and also $\alpha = \tilde{\alpha}$
because partial likelihood
has the same shape
- What if interaction? $\lambda(t|Z) = \lambda_0(t) \cdot \exp\{\alpha \cdot (\text{age} + t) + \beta \cdot Z + \gamma \cdot (\text{age} + t) \cdot Z\}$
 $\tilde{\lambda}(t|Z) = \tilde{\lambda}_0(t) \cdot \exp\{\tilde{\alpha} \cdot (\text{age}) + \tilde{\beta} \cdot Z + \tilde{\gamma} \cdot \text{age} \cdot Z\}$
 $\text{age} = 0 = Z \rightsquigarrow \lambda_0(t) \cdot e^{\alpha t} \text{ vs. } \tilde{\lambda}_0(t)$ and only $\text{age} = 0 \rightsquigarrow \lambda_0(t) \cdot e^{\alpha t + \gamma t Z} \stackrel{Z=0}{=} \lambda_0(t) \cdot e^{\alpha t}$
 under $Z = 0$ we are under the same conditions as above $\rightsquigarrow \alpha$ and $\tilde{\alpha}$ have the same interpretation
- $$\frac{\lambda(t|\text{age}=0, Z)}{\lambda(t|\text{age}=0, Z)} = \exp\{\tilde{\beta}(Z_1 - Z_2)\} \quad \text{but} \quad \frac{\lambda(t|\text{age}=0, Z_1)}{\lambda(t|\text{age}=0, Z_2)} = \frac{\lambda_0(t) \exp\{\alpha t + \beta Z_1 + \gamma t Z_1\}}{\lambda_0(t) \exp\{\alpha t + \beta Z_2 + \gamma t Z_2\}} = \exp\{\beta(Z_1 - Z_2) + \gamma t(Z_1 - Z_2)\}$$

$\lambda(t|\text{age}, Z) = \exp\{\beta + \gamma(\text{age} + t)\}$

generalizing again: $\frac{\lambda(t|\text{age}, Z+1)}{\lambda(t|\text{age}, Z)} = \exp\{\tilde{\beta} + \tilde{\gamma} \cdot \text{age}\}$ but: $\frac{\lambda(t|\text{age}, Z+1)}{\lambda(t|\text{age}, Z)} = \exp\{\beta + \gamma(\text{age} + t)\}$

combined: $\frac{\frac{\lambda(t|\text{age}, Z+1)}{\lambda(t|\text{age}, Z)}}{\frac{\lambda(t|\text{age}, Z+1)}{\lambda(t|\text{age}, Z)}} = e^{\tilde{\gamma}}$

$$\frac{\lambda(t|\text{age}+1, Z+1)}{\lambda(t|\text{age}+1, Z)} = \frac{e^{\beta + \gamma(\text{age} + t + 1)}}{e^{\beta + \gamma(\text{age} + t)}} = e^{\gamma}$$
- $$\text{PL without interaction: } \prod_{i=1}^n \frac{\exp\{\tilde{\alpha} A_i + \tilde{\beta}^T Z_i\}}{\sum_{j=1}^J Y_j(t_i) \exp\{\tilde{\alpha} A_j + \tilde{\beta}^T Z_j\}}$$

$$\text{PL with interaction: } \prod_{i=1}^n \frac{\exp\{\tilde{\alpha} A_i + \tilde{\beta} Z_i + \tilde{\gamma} A_i Z_i\}}{\sum_{j=1}^J Y_j(t_i) \exp\{\tilde{\alpha} A_j + \tilde{\beta} Z_j + \tilde{\gamma} A_j Z_j\}}$$

$\rightsquigarrow \text{in general } (\alpha, \beta, \gamma) \neq (\tilde{\alpha}, \tilde{\beta}, \tilde{\gamma})$
- $$\text{in R: using } \text{tt}(\cdot) \text{ function: } \text{coxph}(\text{Surv}(time, delta) \sim bili + \text{tt}(age), \text{data} = \text{pbc}) \quad (\text{time transformation})$$

in this example yields the same coefficients as $\sim \text{bili} + \text{age}$

however " $\sim \text{bili} * \text{age}$ " and " $\sim \text{bili} * \text{tt}(age)$ " do have similar, but not the same estimates of coefficients
- Warning: time and age need to be in the same units! Ex: $\text{age}[\text{year}] \Rightarrow \text{time}[\text{days}] \sim \text{bili} * \text{tt}(age)$, $\text{tt} = \text{function}(x, t, \dots)$
 $\{x + t / 365.25\}$ within coxph function