

Jan Vávra

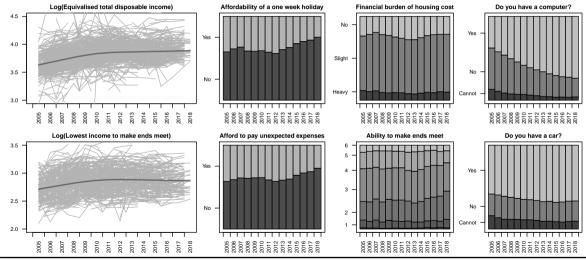
Department of Probability and Mathematical Statistics

GLMM Based Clustering of Multivariate Mixed Type Longitudinal Data

15th June 2022

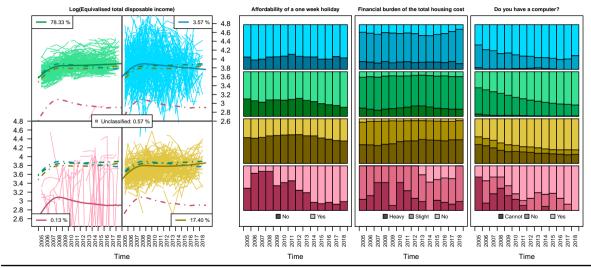
ROBUST 2022

EU-SILC - Outcomes of interest



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Clustering into G = 4 groups



Random effects models

Each outcome is modelled by predictor η consisting of

In case of EU-SILC application:

- Fixed:
 - intercept
 - time quadratic spline with 3 equidistant knots
 - level of urbanisation
 - highest education level achieved in the household
 - presence of a student
 - presence of a baby
 - no interaction terms
- Random: random intercept $\mathbf{Z}_{ii}^{\top}\mathbf{b}_{i}=b_{i,0}\overset{\text{iid}}{\sim}\operatorname{N}\left(0,\sigma_{b}^{2}\right),i=1,\ldots,n$

GLMM for different types of outcomes

Numeric outcome

- Linear Mixed-effects Model (LMM)
 - $Y_{ij}^N \mid \boldsymbol{\eta}^N \sim N(\boldsymbol{\eta}^N, \tau^{-1})$

Ordinal outcome of K levels

- Ordinal Logistic Regression (OLR)
 - $p_k = P[Y^O > k | \eta^O, \mathbf{c}] = logit^{-1} (\eta^O c_k)$
 - $q_k = P[Y^O = k|_{\eta^O}, \mathbf{c}] = p_{k-1} p_k$
 - ordered intercepts

$$-\infty = c_0 < c_1 < \cdots < c_{K-1} < c_K = \infty$$

Binary outcome

Logistic Regression with random effects

•
$$P[Y^B = 1 | \eta^B] = logit^{-1}(\eta^B) = \frac{e^{\eta^B}}{1 + e^{\eta^B}}$$

General categorical outcome of *K* levels

- Multinomial Logistic Regression (MLR)
 - η_k^C specific for each level k

$$P\left[Y^{C} = k | \eta_{1}^{C}, \dots, \eta_{K-1}^{C}\right] =$$

$$= \operatorname{softmax}_{k}(\eta^{C}) = \frac{e^{\eta_{K}^{C}}}{1 + \sum_{k=1}^{K-1} e^{\eta_{K'}^{C}}}$$

Joint modelling

Models for individual outcomes

- Numeric: $Y_i^N \sim \text{LMM}\left(\beta_N, \boldsymbol{b}_i^N, \tau\right)$
- Ordinal: $Y_i^O \sim \mathsf{ORL}\left(\beta_O, \boldsymbol{b}_i^O, \boldsymbol{c}\right)$

- Binary: $Y_i^B \sim \mathsf{LR}\left(\beta_B, \boldsymbol{b}_i^B\right)$
- Categorical: $Y_i^C \sim \text{MLR}\left(\beta_{k,C}, \boldsymbol{b}_i^C\right)$

Join individual models through joint distribution of random effects

$$\boldsymbol{b}_{i} = \begin{pmatrix} \boldsymbol{b}_{i}^{N} \\ \boldsymbol{b}_{i}^{D} \\ \boldsymbol{b}_{i}^{C} \end{pmatrix} \stackrel{\text{iid}}{\sim} N \begin{pmatrix} \boldsymbol{0} = \begin{pmatrix} \boldsymbol{0} \\ \boldsymbol{0} \\ \boldsymbol{0} \\ \boldsymbol{0} \end{pmatrix}, \boldsymbol{\Sigma} = \begin{pmatrix} \boldsymbol{\Sigma}^{NN} & \boldsymbol{\Sigma}^{NB} & \boldsymbol{\Sigma}^{NO} & \boldsymbol{\Sigma}^{NC} \\ \boldsymbol{\Sigma}^{BN} & \boldsymbol{\Sigma}^{BB} & \boldsymbol{\Sigma}^{BO} & \boldsymbol{\Sigma}^{BC} \\ \boldsymbol{\Sigma}^{ON} & \boldsymbol{\Sigma}^{OB} & \boldsymbol{\Sigma}^{OO} & \boldsymbol{\Sigma}^{OC} \\ \boldsymbol{\Sigma}^{CN} & \boldsymbol{\Sigma}^{CB} & \boldsymbol{\Sigma}^{CO} & \boldsymbol{\Sigma}^{CC} \end{pmatrix}, \quad i = 1, \dots, n$$

Joint probability density function

$$\begin{split} h(\boldsymbol{y}_{i};\,\mathcal{C}_{i},\boldsymbol{\zeta}) &= p\left(\mathbb{Y}_{i} = \boldsymbol{y}_{i} \,|\, \boldsymbol{\beta}_{N}, \boldsymbol{\tau}, \boldsymbol{\beta}_{B}, \boldsymbol{\beta}_{O}, \boldsymbol{c}, \boldsymbol{\beta}_{C}, \boldsymbol{\Sigma};\, \mathcal{C}_{i}\right) = \\ &= \int \prod_{r=1}^{R} \prod_{i=1}^{n_{i}} \exp\left\{\ell^{\text{type}(r)}\left(Y_{i,j}^{r} |\, \boldsymbol{b}_{i}^{r}, \boldsymbol{\zeta}_{r};\, \mathcal{C}_{i,j}\right)\right\} \cdot (2\pi)^{-\frac{\dim b_{i}}{2}} \,|\boldsymbol{\Sigma}|^{-1} \exp\left\{-\frac{1}{2}\boldsymbol{b}_{i}^{\top}\boldsymbol{\Sigma}^{-1}\boldsymbol{b}_{i}\right\} \,\mathrm{d}\boldsymbol{b}_{i} \end{split}$$

- Requires methods for numerical evaluation of the integral
- ightarrow solved by Laplacian approximation or generally Adaptive Gaussian Quadrature
 - Pinheiro and Chao (2006)

Model based clustering

To identify different patterns:

- suppose *G* latent groups each following model given by $h(\mathbf{y}_i; C_i, \zeta^{(g)})$
- $\zeta^{(g)}$ consists of
 - \bullet ψ parameters common to all latent groups
 - $\psi^{(g)}$ group-specific parameters (fixed effects for evolution in time, ...)
- group allocation indicators $U_i \in \{1, ..., G\}$
- marginal clustering probabilities $0 < w_g := P[U_i = g] < 1, \quad w_1 + \cdots + w_G = 1$
- ullet complete set of unknown parameters $oldsymbol{ heta} = \{oldsymbol{w}, \psi, \psi^{(1)}, \dots, \psi^{(G)}\}$
- mixture distribution for Y_i:

$$f(\mathbf{y}_i|\mathcal{C}_i;\boldsymbol{\theta}) = \sum_{g=1}^G w_g h\left(\mathbf{y}_i;\mathcal{C}_i,\boldsymbol{\zeta}^{(g)}\right)$$

Clustering probabilities

- Estimation: Bayesian approach and MCMC methods (Gibbs sampling)
- Metropolis proposal steps for fixed and random effects
- By Bayes Theorem:

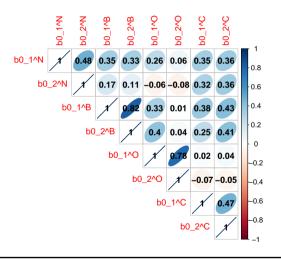
$$p_{i,g}(heta) = \mathsf{P}\left[U_i = g | extbf{Y}_i = extbf{y}_i, \mathcal{C}_i; heta
ight] = rac{w_g h\left(extbf{y}_i; \mathcal{C}_i, extstyle \zeta^{(g)}
ight)}{\sum\limits_{\ell=1}^G w_\ell h\left(extbf{y}_i; \mathcal{C}_i, extstyle \zeta^{(\ell)}
ight)}$$

- → integral approximation required
- Simple clustering rule:

$$\widehat{U}_i := g \iff g = \underset{\ell \in \{1,...,G\}}{\operatorname{arg \, max}} \widehat{\rho_{i,\ell}}$$

- Alternatively use sampled cluster indicators U_i
- Software: implemented in using the C programming language

Correlation matrix of random intercepts



Random intercepts for

Log(Total disposable income)

Log(Lowest income to make ends meet)

Affordability of a one week holiday

Afford to pay unexpected expenses

Ability to make ends meet

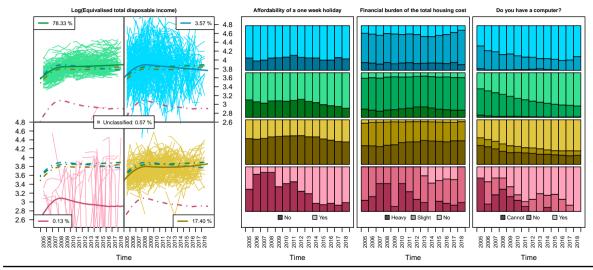
Financial burden of the total housing cost

Do you have a computer?

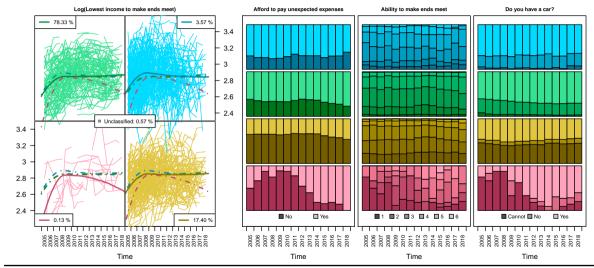
Do you have a car?

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Clustering into $\widehat{\textit{G}}_{+}=4$ groups with respect to time



Clustering into $\widehat{\textit{G}}_{+}=4$ groups with respect to time



Conclusion

See you at my poster (doors to knihovna) with application to PBC data

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Conclusion

See you at my poster (doors to knihovna) with application to PBC data

Thank you for your attention.

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Literature references

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