## Lecture 6 | 15.04.2024

## Parametric structures for variance/covariance in mixed models

## Some overview

$\square$ Consider a linear regression model for repeated measurements within (independent) subjects $i \in\{1, \ldots, N\}$ in a form

$$
\boldsymbol{Y}_{i}=\mathbb{X}_{i} \beta+\varepsilon_{i}
$$

where $\boldsymbol{Y}_{i}=\left(Y_{i 1}, \ldots, Y_{i n_{i}}\right)^{\top}$ is the subject specific response vector, $\mathbb{X}_{i}=\left(\boldsymbol{X}_{i 1}, \ldots, \boldsymbol{X}_{i n_{i}}\right)^{\top}$ for $\boldsymbol{X}_{i j}=\left(X_{i j 1}, \ldots, X_{i j p}\right)^{\top}$ and $j=1, \ldots, n_{i}$ are the subject (and time specific) explanatory vectors (of dimension $p \in \mathbb{N}$ ) and $\boldsymbol{\beta}=\left(\beta_{1}, \ldots, \beta_{p}\right)^{\top}$ is the unknown vector of the regression parameters (mean structure) same for all subjects and time points $\boldsymbol{t}_{i}=\left(t_{i 1}, \ldots, t_{i n_{i}}\right)^{\top}$
$\square$ the variance-covariance structure within each subject is modelled by the vector parameters $\boldsymbol{\alpha} \in \mathbb{R}^{q}$, such that $\varepsilon_{i} \sim N_{n_{i}}\left(\mathbf{0}_{i}, \mathbb{V}_{i}\left(\boldsymbol{t}_{i}, \boldsymbol{\alpha}\right)\right)$
$\square$ the stochastic (non-systematic) term of the model - random errors $\varepsilon_{i j}$ are decomposed into three main parts: the random effects, the serial correlation, and the measurement error

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\varepsilon_{i j}=\boldsymbol{z}_{i j}^{\top} \boldsymbol{w}_{i}+W_{i}\left(t_{i j}\right)+\omega_{i j}
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for random vector $\boldsymbol{w}_{i}$, random process $W_{i}(t)$, and random variable $\omega_{i j}$

## Stochastic properties of the error terms

measurement errors $\omega_{i j} \sim N\left(0, \tau^{2}\right)$, mutually independent for $i$ and $j$
lets denote $\omega_{i}=\left(\omega_{i 1}, \ldots, \omega_{i i_{i}}\right)^{\top}$ and $\omega_{i} \sim N_{n_{i}}\left(\mathbf{0}, \tau^{2} \mathbb{I}_{i}\right)$, for $\mathbb{I}_{i} \in \mathbb{R}^{n_{i} \times n_{i}}$

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$\square$ serial correlation represented by random variables $W_{i}\left(t_{i j}\right)$ sampled from $N \in \mathbb{N}$ independent copies of a stationary Gaussian process $\{W(t) ; t \in \mathbb{R}\}$, with the zero mean, variance $\sigma^{2}>0$ and the correlation function $\rho(u)=\operatorname{cor}(W(t), W(t+u))$
$\square$ random variables $W_{i j} \equiv W_{i}\left(t_{i j}\right)$ are independent with respect to subjects $i \in\{1, \ldots, N\}$ but dependent within subjects, i.e., for indexes $j=1, \ldots, n_{i}$
lets denote $\mathbb{H}_{i}=\left(h_{i j k}\right)_{j, k=1}^{n_{i}}$, where $h_{i j k}=\rho\left(\left|t_{i j}-t_{i k}\right|\right)$, i.e., the correlation between $Y_{i j}$ and $Y_{i k}$, taken at the time points $t_{i j}$ and $t_{i k}$
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$\square$ normally distributed random effects $\boldsymbol{w}_{i} \sim N_{r}(\mathbf{0}, \mathbb{G})$, mutually independent for $i=1, \ldots, N$, with the corresponding explanatory variables $z_{i j} \in \mathbb{R}^{r}$
$\square$ the random effect $\boldsymbol{w}_{i}$ is only subject specific (index $i$ ) but the explanatory vectors $z_{i j}$ related to this random effects are subject and time specific
$\square$ lets denote $\mathbb{Z}_{i}=\left(z_{i 1}, \ldots, z_{i n_{i}}\right)^{\top} \in \mathbb{R}^{n_{i} \times r}$

## Parametric models for variance/covariance

Variance/covariance of $\boldsymbol{Y}_{i} \in \mathbb{R}^{\boldsymbol{n}_{i}}$ can be expressed as
$\square \operatorname{Var} \boldsymbol{Y}_{i}=\operatorname{Var}\left(\varepsilon_{i}\right)=\operatorname{Var}\left[\boldsymbol{z}_{i j}^{\top} \boldsymbol{w}_{i}+W_{i}\left(t_{i j}\right)+\omega_{i j}\right]=\mathbb{Z}_{i} \mathbb{G}_{\mathbb{Z}_{i}^{\top}}^{\top}+\underbrace{\sigma^{2} \mathbb{H}_{i}+\tau^{2} \mathbb{I}_{i}}_{\mathbb{R}_{i} \text { in SAS }}$
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$\hookrightarrow$ because random quantities $\boldsymbol{w}_{i}, \boldsymbol{W}_{i}$, and $\boldsymbol{\omega}_{i}$ are mutually independent
Thus, the mean structure is fully modelled by the specification of the model matrix $\mathbb{X}_{i}$ and the vector of parameters $\beta \in \mathbb{R}^{p}$ but the variance-covariance structure is more complex and it is fully specified by matrices $\mathbb{G}, \mathbb{Z}_{i}$ and $\mathbb{H}_{i}$ and, in addition, two parameters $\sigma^{2}, \tau^{2}>0$
$\square$ As the subjects $i \in\{1, \ldots, N\}$ are independent, we will only investigate different forms for the variance structure in $\operatorname{Var} \boldsymbol{Y}_{i}$, or $\operatorname{Var}\left(\varepsilon_{i}\right)$ respectively, for some generic subject $\boldsymbol{Y} \in \mathbb{R}^{n}$, with $n \in \mathbb{N}$ repeated measurements taken at the time points at $\boldsymbol{t}=\left(t_{1}, \ldots, t_{n}\right)^{\top} \in \mathbb{R}^{n}$
$\square$ the overall variance-covariance structure for $\operatorname{Var} \boldsymbol{Y}_{i}$ will be a block-diagonal matrix with squared matrices of the types $n_{i} \times n_{i}$ in the diagonal

## Example: correlation \& variogram

Consider two different (theoretical) correlation functions $\rho$ and the corresponding variogram functions $\gamma(x)=\sigma^{2}(1-\rho(x))$, for $\sigma^{2}=1$

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## Serial correlation (only) model

$\square$ from the three possible variance/covariance terms in the expansion

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\operatorname{Var} \boldsymbol{Y}=\operatorname{Var}(\varepsilon)=\mathbb{Z} \mathbb{G} \mathbb{Z}^{\top}+\sigma^{2} \mathbb{H}+\tau^{2} \mathbb{I}
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\rho(x)=\exp \left\{-\phi x^{2}\right\}, \phi=1
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$$
\gamma(x)=\sigma^{2}(1-\rho(x)), \sigma^{2}=2
$$



## Parametric models for variance/covariance

Considering both models-the exponential model and the Gaussian correlation model-the role of the $\phi>0$ parameter is the same-as the value of $\phi$ increases, the variograms rises more sharply and the simulated realizations are less smooth
(Fig. 5.1., Diggle et al., 2002)


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(Fig. 5.2., Diggle et al., 2002)


## Ante-dependence model

$\square$ assuming a direct dependence on previous observations in a sense that $\varepsilon_{j}$ taken at the time point $t_{j}>t_{j-1}>t_{j-2}>\ldots$ depends explicitly on some previous $k$ errors $\varepsilon_{j-1}, \ldots, \varepsilon_{j-k} \quad$ (ante-dependence of order $k$ )
$\square$ also well known as the $k$-order Markov model or an autoregressive sequence of order $k, \operatorname{AR}(k)$
$\square \mathrm{AR}(1)$ model formally: $\varepsilon_{j}=\alpha \varepsilon_{j-1}+\omega_{j}$, for $\omega_{j}$ being i.i.d. from $N\left(0, \sigma^{2}\right)$, for $\varepsilon_{0} \sim N\left(0, \sigma^{2} /(1-\alpha)\right)$
$\square$ the ante-dependence model can be problematic for situations with unequally spaced repeated observations within the subject
$\square$ on the other hand, small orders $k \in \mathbb{N}$ can be suitable for the likelihood estimation (straightforward to get the joint distribution/density of $\varepsilon_{i}$ )
$\square$ the ante-dependence property is not preserved when incorporating, for instance, the measurements errors

## Serial correlation + measurement errors model

$\square$ models where there are no random effects present and the variance of $\varepsilon_{i}$ reduces to

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\operatorname{Var}(\varepsilon)=\sigma^{2} \mathbb{H}+\tau^{2} \mathbb{I}
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$\square$ the variance of $\varepsilon_{j}$ is captured by the sum $\tau^{2}+\sigma^{2}$ with the corresponding variogram function $\gamma(u)=\tau^{2}+\sigma^{2}(1-\rho(u))$
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$\gamma(x)=\sigma^{2}(1-\rho(x)), \sigma^{2}=2, \tau^{2}=0.5$


## Random intercept model

$\square$ the simplest example of a general model with three variance/covariance terms in the decomposition where $z_{i j}=1$ and $w_{i} \sim N\left(0, \nu^{2}\right)$
$\square$ the variance of $\varepsilon_{j}$ is $\operatorname{Var} \varepsilon_{j}=\nu^{2}+\sigma^{2}+\tau^{2}$ and the corelation within the whole vector $\varepsilon$ is captured by matrices $\mathbb{H}$ and $\mathbb{J}$
(Diggle, 1988)

- $\operatorname{Var}(\varepsilon)=\nu^{2} \mathbb{J}+\sigma^{2} \mathbb{H}+\tau^{2} \mathbb{I}$ with the correlation function $\rho(x)$ and the variogram function $\gamma(u)=\tau^{2}+\sigma^{2}(1-\rho(u))$
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$$
\rho(x)=\exp \left\{-\phi x^{2}\right\}, \phi=1
$$


$\gamma(x), \sigma^{2}=2, \tau^{2}=0.5, \nu^{2}=0.3$


## Random intercept and slope model

$\square$ more general model allows for some form of nonstationarity-the variability within the subject now depends on the time
$\square$ the random effects are now $\boldsymbol{w}_{i} \sim N_{2}(\mathbf{0}, \mathbb{G})$ and for simplicity we assume that $\mathbb{G}=\nu^{2} \mathbb{I}$, the covariates are $\boldsymbol{z}_{i j}=\left(1, t_{i j}\right)^{\top}$
$\square$ the variance of $\varepsilon_{j}$ is $\nu^{2}\left(1+t_{j}^{2}\right)+\sigma^{2}+\tau^{2}$ and the whole variance/covariance structure of $\varepsilon$ is Var $\varepsilon=\mathbb{Z} \mathbb{G} \mathbb{Z}^{\top}+\sigma^{2} \mathbb{H}+\tau^{2} \mathbb{I}$, where the matrix $\mathbb{Z} \mathbb{G} \mathbb{Z}^{\top}$ is a $n \times n$ matrix with elements $t_{j} t_{k}$, for $j, k=1, \ldots, n$

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how can by this structure revealed with the sample variogram?

$$
\widehat{\gamma}_{i}(u)=\frac{1}{2(n-u)} \sum_{j=u+1}^{n}\left[\varepsilon_{i j}-\varepsilon_{i(j-u)}\right]^{2}, \quad \text { for } u \in\{0, \ldots, n-1\}
$$

## Random effects and measurement error

$\square$ serial correlation appear to be quite natural feature for the longitudinal data analysis but in some applications the effects of the serial correlation may be dominated by the random effects and measurement errors
$\square$ particularly, if $\sigma^{2}>0$ is small compared to $\nu^{2}+\tau^{2}$
$\square$ the model error $\varepsilon_{i j}$ reduces to $\varepsilon \boldsymbol{Z}_{i j}^{\top} \boldsymbol{w}_{i}+\omega_{i j}$ and the corresponding variance/covariance structure is of the form Var $\varepsilon=\mathbb{Z} \mathbb{G} \mathbb{Z}^{\top}+\tau^{2} \mathbb{I}$
$\square$ the simplest scenario involves the random intercept only, meaning that $\mathbb{Z} \mathbb{G} \mathbb{Z}$ reduces to the matrix $\mathbb{J}$ and $\operatorname{Var} \varepsilon=\nu^{2} \mathbb{J}+\tau^{2} \mathbb{I}$
$\square$ also called a "split-plot" model (because it is equivalent with the correlation induced by the randomization for a classical split-plot experiment)

## Model selection \& model building

$\square$ Practical utilization of the model
firstly, it is important to be able to validly answer the question of interest
statistical inference-statistical tests and confidence intervals/regions
$\square$ Conditional mean structure
$\square$ exploratory in terms of some visualization tools (plots, graphs, etc.)
$\square$ modeling in terms of the model matrix $\mathbb{X}$
$\square$ Designe of experiment
$\square$ many existing problems could be avoided by a proper experiment planning
$\square$ balanced data, proper randomization, treatment assignments, etc.
$\square$ Variace/covariance structure
$\square$ exploratory in terms of some residuals inspection
$\square$ effects of unobserved covariates,

## Different covariance structures in SAS

See, for instance, the implementation of PROC MIXED in SAS and the corresponding SAS help/tutorial
$\square$ variance components
$\square$ compound symmetry
$\square$ unstructured
$\square$ autoregressive
$\square$ spatial
$\square$...

SAS Documentation at https://documentation.sas.com

## Model diagnostics

The main idea of the statistical modeling process:
model formulation $\rightarrow$ model estimation $\rightarrow$ statistical inference $\rightarrow$ model diagnostics

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model formulation $\rightarrow$ model estimation $\rightarrow$ statistical inference $\rightarrow$ model diagnostics
$\square$ The mean structure
$\square$ simple empirical characteristics, data scatterplots
$\square$ simple summary plots/graphs (e.g., boxplots)
$\square$...
$\square$ The variance/covariance structure
$\square$ residual inspection and various residual plots
$\square$ sample variogram function and its alternatives
$\square$ empirical variogram vs. fitted variogram (model based)
$\square$...

## Summary

$\square$ exploratory analysis
$\square$ model building
$\square$ model diagnostics
$\square$ confirmatory analysis
$\square$ model interpretation

