

Lecture 7 | 09.04.2024

Statistical inference

in a normal linear model

Overview

- In general, the random sample $\{(Y_i, \mathbf{X}_i^\top)^\top; i = 1, \dots, n\}$ from some joint distribution $F_{(Y, \mathbf{X})}$ (a generic random vector $(Y, \mathbf{X}^\top)^\top \in \mathbb{R}^{p+1}$)
- the underlying structure (i.e., linear model) is assumed to hold

$$Y_i = \mathbf{X}_i^\top \boldsymbol{\beta} + \varepsilon_i, \quad \text{for } i = 1, \dots, n, \text{ where } \varepsilon_i \sim N(0, \sigma^2)$$

- the model can be equivalently also expressed in a matrix notation as

$$\mathbf{Y} = \mathbb{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}, \quad \text{where } \boldsymbol{\varepsilon} \sim N_n(\mathbf{0}, \sigma^2\mathbb{I})$$

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- the joint distribution function $F_{Y, \mathbf{X}}(y, \mathbf{x})$ can be factorized as

$$F_{Y, \mathbf{X}}(y, \mathbf{x}) = F_{Y|\mathbf{X}}(y|\mathbf{x}) \cdot F_{\mathbf{X}}(\mathbf{x})$$

and $F_{Y|\mathbf{X}}(y|\mathbf{x})$ is assumed to be a conditional normal distribution

Typical linear model assumptions

□ Ordinary linear regression model

- random sample (Y_i, \mathbf{X}_i) for $i = 1, \dots, n$ from the joint distribution $F_{(Y, \mathbf{X})}$
- mean specification $E[\mathbf{Y}|\mathbb{X}] = \mathbb{X}\beta$, respectively $E[Y|\mathbf{X}] = \mathbf{X}^\top \beta$
- variance specification $\text{Var}(\mathbf{Y}|\mathbb{X}) = \sigma^2 \mathbb{I}$, resp. $\text{Var}(\varepsilon) = \sigma^2 \mathbb{I}$

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The formulation of the normal model above also implies the following:

- $\varepsilon|\mathbb{X} \sim N_n(\mathbf{0}, \sigma^2 \mathbb{I})$
- $\varepsilon \sim N_n(\mathbf{0}, \sigma^2 \mathbb{I})$
- error terms $\varepsilon_1, \dots, \varepsilon_n$ form a random sample from a univariate normal distribution with the zero mean and the variance $\sigma^2 > 0$

Parameter estimation in the normal model

- the normal model $\mathbf{Y} = \mathbb{X}\beta + \varepsilon$, where $\varepsilon \sim N_n(\mathbf{0}, \sigma^2\mathbb{I})$ is assumed to hold
- the unknown parameters to be estimated are $\beta \in \mathbb{R}^p$, and $\sigma^2 > 0$
- statistical inference (confidence intervals, statistical tests) could be, however, also performed with respect to the parameters β and σ^2 but, it can be also of some interest to do inference about some linear combination(s) of β
- from the practical point of view, we are interested in the parameter vector β itself but also some (reasonable) linear combinations $\mathbf{l}^\top\beta$, for some (fixed) vector $\mathbf{l} \in \mathbb{R}^p$

Parameter estimation in a normal model

There are basically two standard techniques for the parameter estimation under the linear model formulation:

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In both situations the estimates are given by the formulae

- ❑ $\hat{\beta} = (\mathbb{X}^T \mathbb{X})^{-1} \mathbb{X} \mathbf{Y}$, where $\mathbb{X}^T \mathbb{X}$ is of a full rank $p \in \mathbb{N}$
- ❑ $\hat{\sigma}^2 = \frac{1}{n-p} \sum_{i=1}^n (Y_i - \hat{Y}_i)^2$, where $\hat{Y}_i = \mathbf{X}_i^T \hat{\beta}$

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Both estimates—quantities $\hat{\beta}$ and $\hat{\sigma}^2$ —are random quantities (random vector and random variable) and, therefore, it is reasonable to investigate their statistical properties (e.g., mean, variance, distribution, etc.)

Linear combinations of the model parameters

- the unknown vector of parameters $\beta \in \mathbb{R}^p$ is used to model the conditional mean structure $E[\mathbf{Y}|\mathbf{X}]$ but specific interpretation (meaning) of the elements of β depends on the parametrization that is used in the model
- therefore, it is also of some interest to perform statistical inference about some linear combination of the unknown vector of parameters—inference about some different parametrization of the mean structure
- let $\mathbf{L} \in \mathbb{R}^{m \times p}$ be a matrix with nonzero rows $\mathbf{l}_1^\top, \dots, \mathbf{l}_m^\top$ and let $\theta = \mathbf{L}\beta = (\mathbf{l}_1^\top \beta, \dots, \mathbf{l}_m^\top \beta)^\top = (\theta_1, \dots, \theta_m) \in \mathbb{R}^m$ be some linear combinations of the original β vector
- instead of performing the inference about $\beta \in \mathbb{R}^p$ we can be interested in performing the statistical inference about $\theta \in \mathbb{R}^m$

Statistical properties of $\hat{\beta}$ and $\hat{\theta}$

Recall, that we are working with the normal linear model of the form $\mathbf{Y}|\mathbb{X} \sim N_n(\mathbb{X}\beta, \sigma^2\mathbb{I})$ and $\hat{\beta} = (\mathbb{X}^\top \mathbb{X})^{-1} \mathbb{X}^\top \mathbf{Y}$ is the estimate for $\beta \in \mathbb{R}^p$

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- $\frac{1}{m}(\hat{\theta} - \theta)^\top (MSe \cdot \mathbb{V})^{-1} (\hat{\theta} - \theta) \sim F_{m, n-p}$, provided that $rank(\mathbb{L}) = m \leq p$

Inference in a normal linear model

□ Inference about some β_j

- confidence interval $\hat{\beta}_j \pm t_{n-p}(1 - \alpha/2)\sqrt{MSe \cdot v_{jj}}$, where $\text{Var}\hat{\beta}_j = \sigma^2 v_{jj}$
- statistical tests of the null hypothesis $H_0 : \beta_j = \beta_j^{(0)}$

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□ Simultaneous confidence region for β

- $S(\alpha) = \{\beta \in \mathbb{R}^p; \frac{1}{p}(\beta - \hat{\beta})^\top (MSe^{-1} \mathbb{X}^\top \mathbb{X})(\beta - \hat{\beta}) < F_{p, n-p}(1 - \alpha)\}$,
which is an ellipsoid with the center $\hat{\beta}$, the shape matrix $MSe \cdot (\mathbb{X}^\top \mathbb{X}^{-1})$
and the diameter $\sqrt{kF_{p, n-p}(1 - \alpha)}$
- statistical test of the null hypothesis $H_0 : \beta = \beta^{(0)}$

Model based predictions

□ Model utilization for

- characterization of the conditional distribution of Y given \mathbf{X}
- explaining the effect of some covariate X_j on the variable Y
- prediction of a new observation Y_{new} given the observed value of \mathbf{X}_{new}

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- ❑ straightforward prediction in terms of the estimated conditional expectation $\hat{\mu}_{new} = \mathbf{X}_{new}^T \hat{\beta}$
- ❑ however, can we do better (e.g., accounting for the variability in Y_{new})?

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- however, can we do better (e.g., accounting for the variability in Y_{new})?
- **distributional assumption**

$$Y_{new} | \mathbf{X}_{new} \sim N(\mathbf{X}_{new}^\top \beta, \sigma^2)$$

where $(Y_{new}, \mathbf{X}_{new})$ is independent of $\{(Y_i, \mathbf{X}_i); i = 1, \dots, n\}$

Theoretical background of the prediction

□ **Formally**

$$Y_{new} = \mathbf{X}_{new}^T \boldsymbol{\beta} + \varepsilon_{new}, \quad \text{for } \varepsilon_{new} \sim N(0, \sigma^2)$$

□ **Theoretical property**

$$P[Y_{new} \in (\mathbb{X}_{new}^T \boldsymbol{\beta} \pm u_{1-\alpha/2} \sigma)] = 1 - \alpha$$

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$$P\left[Y_{new} \in (\mathbb{X}_{new}^T \hat{\boldsymbol{\beta}} \pm t_{1-\alpha/2}(n-p) \sqrt{1 + \mathbf{X}_{new}^T (\mathbb{X}^T \mathbb{X})^{-1} \mathbf{X}_{new}})\right] = 1 - \alpha$$

Summary

□ Simple inference in a normal linear model

- confidence intervals and statistical tests for elements of $\beta \in \mathbb{R}^p$
- confidence intervals for some linear combination of $I^\top \beta$, for $I \in \mathbb{R}^p$

□ Simultaneous inference for vector parameters in the linear model

- confidence regions and statistical tests for the whole vector $\beta \in \mathbb{R}^p$
- confidence regions for some linear combinations $L\beta$, where $L \in \mathbb{R}^{m \times p}$

□ Prediction in the normal linear model

- point prediction for a new value of Y given the observed $\mathbf{X} = \mathbf{x}$
- interval prediction for a new value of Y given the observed $\mathbf{X} = \mathbf{x}$