

# Testy pre regresné kvantily založené na metóde sedlového bodu

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# Obsah

- 1 Úvod
- 2 Testy pre regresné kvantily
- 3 Simulácie

# Metóda sedlového bodu

- aproximácie hustoty odhadov metódou sedlového bodu predstavil Daniels (1954), neskôr rozpracoval Hampel (1973), Field a Hampel (1982)
- veľmi presné aproximácie hustoty odhadov
- testy pre M-odhady založené na výraze v exponente odhadu hustoty

## Model

model

$$Y_i = \mathbf{X}_i^T \boldsymbol{\beta} + u_i, \quad i = 1, \dots, n,$$

kde  $u_i \sim \frac{1}{\sigma} g(\cdot)$  a  $(Y_i, \mathbf{X}_i)$  sú i.i.d. s hustotou  $\frac{1}{\sigma} g\left(\frac{y_i - \mathbf{x}_i^T \boldsymbol{\beta}}{\sigma}\right) k(\mathbf{x}_i)$ ,  $\boldsymbol{\beta} \in \mathbb{R}^{p+1}$

- regresný kvantil je riešením

$$\boldsymbol{\beta}_\alpha = \arg \min \{ \mathbb{E} \rho_\alpha(Y - \mathbf{X}\mathbf{t}) : \mathbf{t} \in \mathbb{R}^p \},$$

kde

$$\rho_\alpha(x) = |x| \{ (1 - \alpha) \mathbf{I}[x < 0] + \alpha \mathbf{I}[x > 0] \}$$

- odhad regresného kvantilu  $\hat{\boldsymbol{\beta}}_\alpha$  je riešením

$$\arg \min_{\mathbf{t}} \sum_{i=1}^n (\alpha - \mathbf{I}[Y_i - \mathbf{X}_i^T \mathbf{t} < 0]) \mathbf{X}_i$$

- $\hat{\boldsymbol{\beta}}_\alpha$  je konzistentným odhadom  $(\beta_0 + \mathbf{G}^{-1}(\alpha), \beta_1, \dots, \beta_p)$

Testujeme hypotézu

$$H_0 : \boldsymbol{\beta}_\alpha = \boldsymbol{\beta}_{\alpha 0}$$

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## Parametrický test

kumulatívna vytvárajúca funkcia skórov  $\psi(Y_i, \beta_\alpha)$ 

$$K_\psi(\lambda, \beta_\alpha) = \log \int \left\{ e^{\alpha \lambda^T \mathbf{x}_i} k(\mathbf{x}_i) \left( e^{-\lambda^T \mathbf{x}_i} G\left(\frac{\mathbf{x}_i^T (\beta_\alpha - \beta)}{\sigma}\right) + 1 - G\left(\frac{\mathbf{x}_i^T (\beta_\alpha - \beta)}{\sigma}\right) \right) \right\} d\mathbf{x}_i$$

testová štatistika

$$h(\hat{\beta}_\alpha) = \log \int \left( \frac{G\left(\frac{\mathbf{x}_i^T (\hat{\beta}_\alpha - \beta_0)}{\sigma}\right)}{\alpha} \right)^\alpha \left( \frac{1 - G\left(\frac{\mathbf{x}_i^T (\hat{\beta}_\alpha - \beta_0)}{\sigma}\right)}{1 - \alpha} \right)^{1-\alpha} k(\mathbf{x}_i) d\mathbf{x}_i$$

za platnosti  $H_0$ 

$$2nh(\hat{\beta}_\alpha) \xrightarrow[n \rightarrow \infty]{\mathcal{D}} \chi_{p+1}^2$$

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## Neparametrický test

- reziduá a ich empirická distribučná funkcia

$$\begin{aligned}r_i &= y_i - \mathbf{x}_i^T \hat{\beta}_\alpha \\l_i &= I[r_i < 0] \\l_{ij}(\beta) &= I[r_i + \mathbf{x}_j^T (\hat{\beta}_\alpha - \beta) < 0] \\G_n^j(\beta) &= \frac{1}{n} \sum_{i=1}^n l_{ij}(\beta)\end{aligned}$$

- empirická hypotetická distribučná funkcia

$$W_{ij} = \frac{\left( \frac{\alpha}{1-\alpha} \frac{1-G_n^j(\beta_{\alpha 0})}{G_n^j(\beta_{\alpha 0})} \right)^{l_{ij}(\beta_{\alpha 0})}}{\sum_{k=1}^n \left( \frac{\alpha}{1-\alpha} \frac{1-G_n^k(\beta_{\alpha 0})}{G_n^k(\beta_{\alpha 0})} \right)^{l_{kj}(\beta_{\alpha 0})}}, i, j = 1, \dots, n$$



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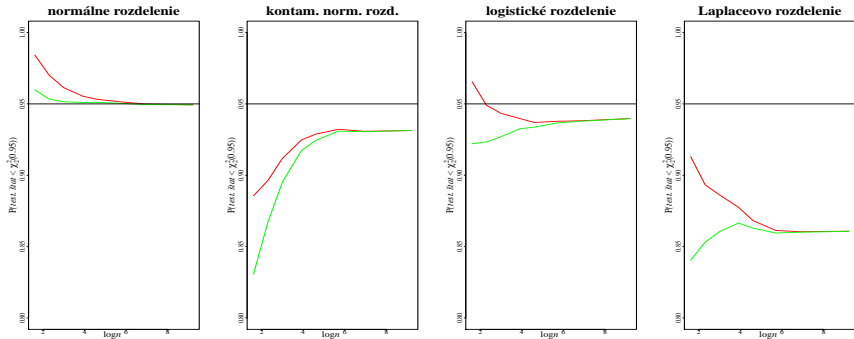
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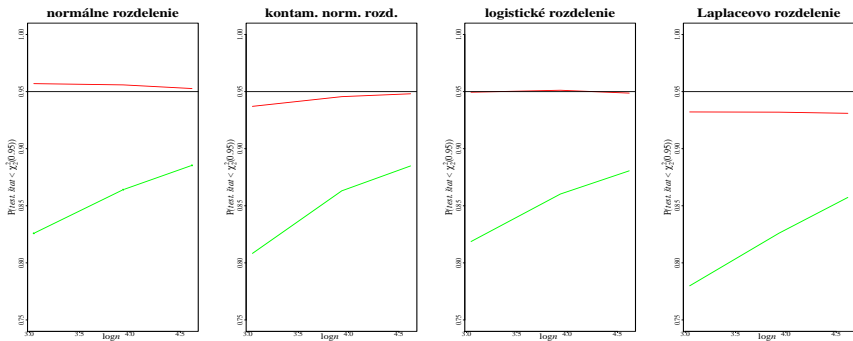
Potom za platnosti  $H_0$

$$-2 \sum_{j=1}^n \log \left\{ \left( \frac{\sum_{i=1}^n w_{ij} l_i}{\alpha} \right)^\alpha \left( \frac{\sum_{i=1}^n w_{ij} (1 - l_i)}{1 - \alpha} \right)^{1-\alpha} \right\} \xrightarrow[n \rightarrow \infty]{\mathcal{D}} \chi_{p+1}^2$$






# Parametrický test - výsledky



# Neparametrický test - výsledky



## Literatúra

-  Daniels, HE  
Saddlepoint approximations on statistics  
Ann. Math. Statist. 25, 631–650.
-  Field, CA and Hampel, FR (1982)  
Small-sample asymptotic distribution of M-estimators of location.  
Biometrika 69, 29–46.
-  Hampel, F. R. (1973). Some small-sample asymptotics. In: Proc. Prague Symposium on Asymptotic Statistics (J. Hájek, ed.), pp. 109–126. Charles University in Prague.
-  Ronchetti E, Robinson J, Young GA (2003)  
Saddlepoint Approximations and Tests Based on Multivariate M-estimates  
The Annals of Statistics 31: 1154–1169
-  Ronchetti E, Welsh AH (1994)  
Empirical Saddlepoint Approximations for Multivariate M-estimators  
Journal of the Royal Statistical Society 56, no.2: 313–326