

Simultánne testovanie strednej hodnoty a variancie normálneho rozdelenia

Martina Chvosteková

Ústav merania,
Slovenská akadémia vied

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Obsah

- 1 Úvod
 - $H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2$ proti $H_1 : \mu \neq \mu_0 \vee \sigma^2 \neq \sigma_0^2$
 - Exatné testy a oblastí spoľahlivosti
- 2 Porovnanie testov a oblastí spoľahlivosti
 - Plochy oblastí
 - Sila testu
- 3 $H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2$ proti $H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$
 - Porovnanie: Silofunkcie testov, obsahy oblastí spoľahlivosti
- 4 Porovnanie testov a oblastí spoľahlivosti
 - Plochy oblastí

$$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2 \text{ proti } H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$$

Porovnanie testov a oblastí spoľahlivosti
Porovnanie testov a oblastí spoľahlivosti

X_1, X_2, \dots, X_n je náhodný výber z normálneho rozdelenia so neznámou strednou hodnotou μ a s neznámou varianciou σ^2

- ozn. *parametrický priestor*

$$\Theta = \{\theta = (\mu, \sigma^2) : -\infty < \mu < \infty, \sigma^2 > 0\}$$

Simultánne testovanie strednej hodnoty a variancie $N(\mu, \sigma^2)$

$$H_0 : (\mu, \sigma^2) = (\mu_0, \sigma_0^2) \in \Theta_0 \text{ proti } H_1 : (\mu, \sigma^2) \in \Theta - \Theta_0 = \Theta_1$$

- hladina významnosti testu $\alpha \in (0, 1)$

$$\bar{X} \sim N(\mu_0, \sigma_0^2/n), \quad \frac{(n-1)S^2}{\sigma_0^2} \sim \chi_{n-1}^2$$

$$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2 \text{ proti } H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$$

Porovnanie testov a oblastí spoľahlivosti
Porovnanie testov a oblastí spoľahlivosti

Prehľad literatúry

Mood, A. M.: Introduction to the Theory of Statistics, *New York, McGraw-Hill*, 1950.

Wilks, S. S.: Mathematical Statistics, *New York, John Wiley and Sons*, 1962.

Arnold, B.C. - Shavelle, R. M.: Joint confidence sets for the mean and variance of a normal distribution, *The American Statistician* 52, 1998, 133–140.

Choudhari, P. - Kundu, D. - Misra, N.: Likelihood ratio test for simultaneous testing of the mean and variance of a normal distribution, *Journal of Statistical Computation and Simulation* 71, 2001, 313–333.

$$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2 \text{ proti } H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$$

Porovnanie testov a oblastí spoľahlivosti
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Arnold a Shavelle (1998)

LSR1 (Large Sample Region)

$$\mathcal{R} = \left\{ (\mu, \sigma^2) : \frac{n(\bar{X} - \mu)^2}{\sigma^2} + \frac{n}{2\sigma^4} (S^2 - \sigma^2)^2 \leq \chi_2^2(1 - \alpha) \right\}$$

LSR2

$$\mathcal{R} = \left\{ (\mu, \sigma^2) : \frac{n(\bar{X} - \mu)^2}{S^2} + \frac{n}{2S^4} (S^2 - \sigma^2)^2 \leq \chi_2^2(1 - \alpha)^* \right\}$$

$$*2F_{2,n-2}(1 - \alpha)$$

Asymptotické rozdelenie

$$\sqrt{n}(S^2 - \sigma_0^2) \approx N(0, 2\sigma_0^2)$$

Modifikovaný Arnold a Shavelle a navrhnutá LSR3

$$\mathcal{R}_{LSR1} = \left\{ (\mu, \sigma^2) : \frac{n(\bar{X} - \mu)^2}{\sigma^2} + \frac{n}{2\sigma^4} (S^2 - \sigma^2)^2 \leq c_{21}(1 - \alpha) \right\}$$

$$\mathcal{R}_{LSR2} = \left\{ (\mu, \sigma^2) : \frac{n(\bar{X} - \mu)^2}{S^2} + \frac{n}{2S^4} (S^2 - \sigma^2)^2 \leq c_{22}(1 - \alpha) \right\}$$

$$\mathcal{R}_{LSR3} = \left\{ (\mu, \sigma^2) : \frac{n(\bar{X} - \mu)^2}{\sigma^2} + \left(\sqrt{\frac{2(n-1)S^2}{\sigma^2}} - \sqrt{2v-1} \right)^2 \leq c_{23}(1 - \alpha) \right\}$$

Asymptotické rozdelenie¹

$$\sqrt{2\chi_v^2} \approx N(\sqrt{2v-1}, 1)$$

¹Fisher, R. A. (1928), *Statistical Methods for Research Workers*, 2nd Edition, pp. 96-97.

$$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2 \text{ proti } H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$$

Porovnanie testov a oblastí spoľahlivosti
Porovnanie testov a oblastí spoľahlivosti

Mood, Wilks a LRT

Mood (1950)

$$\mathcal{R} = \{(\mu, \sigma^2) : \sqrt{n}|\bar{X} - \mu|/\sigma \leq u(1 - \alpha_1/2) \\ \wedge \chi_{n-1}^2(\alpha_2/2) \leq (n-1)S^2/\sigma^2 \leq \chi_{n-1}^2(1 - \alpha_2/2)\}$$

$$- 1 - \alpha = (1 - \alpha_1)(1 - \alpha_2) \text{ a } \alpha_1 = \alpha_2 = 1 - \sqrt{1 - \alpha}$$

Wilks (1962)

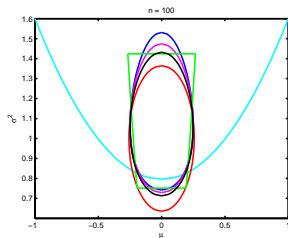
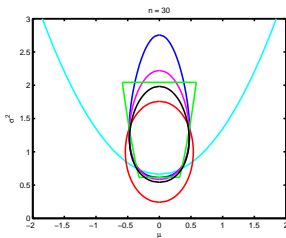
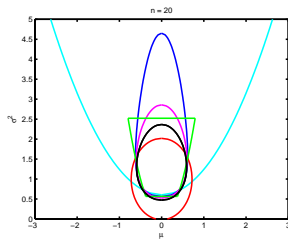
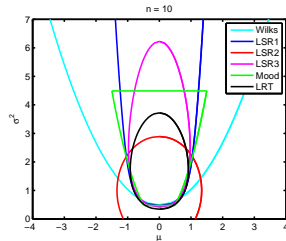
$$\mathcal{R} = \left\{ (\mu, \sigma^2) : \frac{n(\bar{X} - \mu)^2}{\sigma^2} + \frac{(n-1)S^2}{\sigma^2} \leq \chi_n^2(1 - \alpha) \right\}$$

Choudhari–Kundu–Misra (2001) - LRT

$$\mathcal{R} = \left\{ (\mu, \sigma^2) : \frac{n(\bar{X} - \mu)^2}{\sigma^2} + \frac{(n-1)S^2}{\sigma^2} - n \ln \left(\frac{(n-1)S^2}{n\sigma^2} \right) - n \leq d_n(1 - \alpha) \right\}$$

$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2$ proti $H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$
 Porovnanie testov a oblastí spoľahlivosti
 Porovnanie testov a oblastí spoľahlivosti

Tvary oblastí spoľahlivosti pre $\alpha = 0,05$ a rôzne rozsahy výberu



$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2$ proti $H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$
 Porovnanie testov a oblastí spoľahlivosti
 Porovnanie testov a oblastí spoľahlivosti

Porovnávacia štúdia

Porovnávame:

- obsahy: LSR1, LSR2, LSR3, Mood, LRT

(pre $\alpha \in \{0,10; 0,05; 0,01\}$ a $n \in \{10, 20, 30, 100\}$)

obsah = [] $\times S^3$

- sily testov: Wilks, LSR1, LSR2, LSR3, Mood, LRT

(pre $\alpha = 0,05$ a $n \in \{10, 20, 30, 100\}$)

Silofunkcie testov závisia od $(\mu_1 - \mu_0)^2$, σ_1^2/σ_0^2

- uvažovali sme $\mu_0 = 0$ a $\sigma_0^2 = 1$
- meníme μ_1 pri pevnom $\sigma_1^2 = 1$
- meníme σ_1^2 pri pevnom $\mu_1 = 0$
- meníme μ_1 a zároveň meníme aj σ_1^2

Obsahy $(1 - \alpha)$ -spoľahlivých oblastí

Metóda	LSR1	LSR2	LSR3	Mood	Mood*	LRT
n	$\alpha = 0, 10$					
10	36,1507	4,1982	5,1218	4,6742	3,2702	3,2079
20	2,0077	1,4259	1,5242	1,5412	1,3104	1,2616
30	1,0185	0,8402	0,8802	0,9059	0,8088	0,7819
100	0,2275	0,2158	0,2200	0,2300	0,2224	0,2129

Mood*

$$\mathcal{R} = \{(\mu, \sigma^2) : \sqrt{n}|\bar{X} - \mu|/\sigma \leq u(1 - \alpha_1/2) \\ \wedge \chi_{n-1}^2(\delta) \leq (n-1)S^2/\sigma^2 \leq \chi_{n-1}^2(1 - \alpha_2 + \delta)\}$$

$$H_0: \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2 \text{ proti } H_1: \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$$

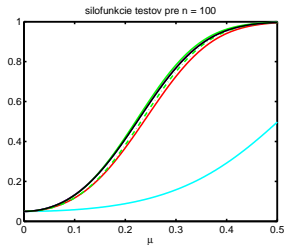
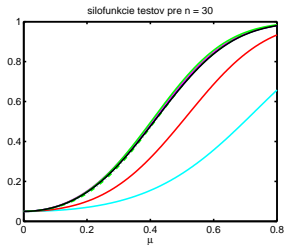
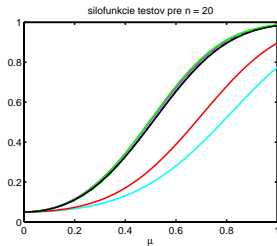
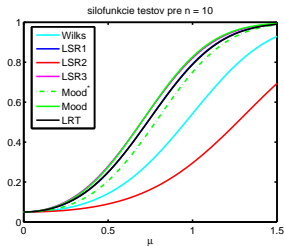
Porovnanie testov a oblastí spoľahlivosti
Porovnanie testov a oblastí spoľahlivosti

Obsahy $(1 - \alpha)$ -spoľahlivých oblastí

Metóda	LSR1	LSR2	LSR3	Mood	Mood*	LRT
n	$\alpha = 0,05$					
10	-	6,4583	8,8844	7,3943	5,1747	4,8910
20	3,8435	2,2946	2,2086	2,2091	1,9069	1,766
30	1,5887	1,2771	1,2236	1,2606	1,1291	1,0668
100	0,3075	0,2949	0,2914	0,3078	0,2995	0,2808
	$\alpha = 0,01$					
10	-	17,881	32,57	18,028	12,617	10,959
20	-	4,8945	4,4771	4,2564	3,5908	3,224
30	4,8396	2,7298	2,2175	2,263	2,0225	1,8329
100	0,5265	0,5246	0,4676	0,5034	0,4868	0,4458

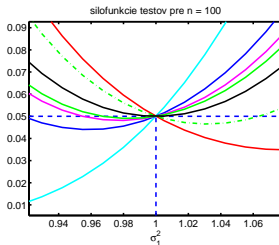
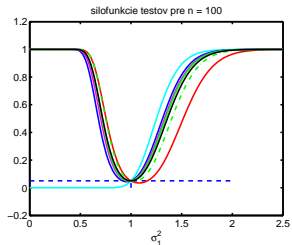
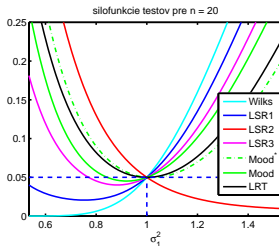
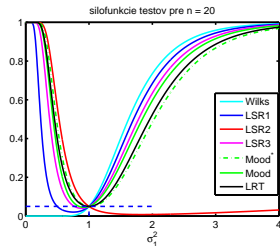
$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2$ proti $H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$
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Silofunkcie pre rôzne μ , ak $\alpha = 0,05$; $\sigma_0^2 = \sigma_1^2 = 1$



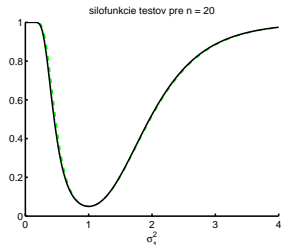
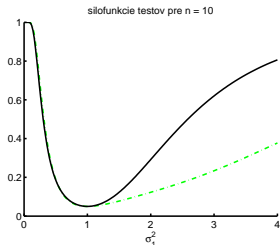
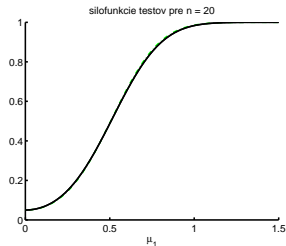
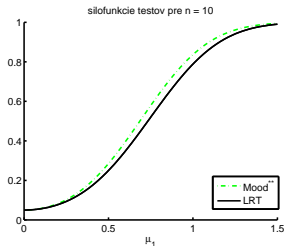
Porovnanie testov a oblastí spoľahlivosti
 $H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2$ proti $H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$
 Porovnanie testov a oblastí spoľahlivosti

Silofunkcie pre rôzne σ_1^2 , ak $\alpha = 0,05$; $\mu_0 = \mu_1 = 0$



$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2$ proti $H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$
 Porovnanie testov a oblastí spoľahlivosti
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Silofunkcie LRT a Mood** , $\alpha = 0.05$



Mood ('optimálne') $\alpha_1, \alpha_2, \delta$ pre $n = 20$

	$\alpha = 0,10$			$\alpha = 0,05$		
	α_1	α_2	δ	α_1	α_2	δ
Mood	,0513	,0513	,0257	,0253	,0253	,0127
Mood*	,0428	,0598	,0532	,0168	,0338	,0325
Mood**	,0446	,0580	,0447	,0215	,0291	,0231

$$1 - \alpha = (1 - \alpha_1)(1 - \alpha_2)$$

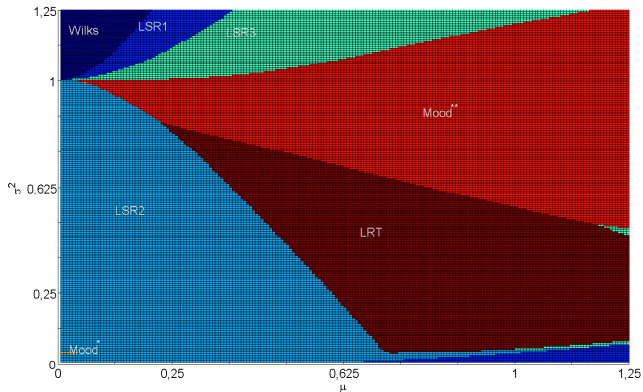
$$\text{Mood} \quad \alpha_1 = \alpha_2 = 1 - \sqrt{1 - \alpha}$$

Mood*, Mood**

$$\mathcal{R} = \{(\mu, \sigma^2) : \sqrt{n}|\bar{X} - \mu|/\sigma \leq u(1 - \alpha_1/2) \\ \wedge \chi_{n-1}^2(\delta) \leq (n-1)S^2/\sigma^2 \leq \chi_{n-1}^2(1 - \alpha_2 + \delta)\}$$

Porovnanie testov a oblastí spoľahlivosti
 $H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2$ proti $H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$
 Porovnanie testov a oblastí spoľahlivosti

Silofunkcie pre rôzne μ_1, σ_1^2 pre $\alpha = 0,05; n = 10$



Iná alternatíva

X_1, X_2, \dots, X_n je náhodný výber z normálneho rozdelenia so neznámou strednou hodnotou μ a s neznámou varianciou σ^2

$$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2 \text{ proti } H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$$

Presné testy, presné oblasti spoľahlivosti

Upravený Mood

$$\mathcal{R} = \left\{ (\mu, \sigma^2) : \sqrt{n}|\bar{X} - \mu|/\sigma \leq u(1 - \alpha_1/2) \right. \\ \left. \wedge (n-1)S^2/\sigma^2 \geq \chi_{n-1}^2(\alpha_2) \right\}$$

$$- 1 - \alpha = (1 - \alpha_1)(1 - \alpha_2) \text{ a } \alpha_1 = \alpha_2 = 1 - \sqrt{1 - \alpha}$$

Návrh upraviť LRT

$$\mathcal{R} = \left\{ (\mu, \sigma^2) : \frac{n(\bar{X} - \mu)^2}{\sigma^2} + \frac{(n-1)S^2}{\sigma^2} - n \ln \left(\frac{(n-1)S^2}{n\sigma^2} \right) - n \leq d_n(1 - \alpha) \right\}$$

Presné testy, presné oblasti spoľahlivosti

Upravený Mood

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$$\mathcal{R} = \left\{ (\mu, \sigma^2) : \frac{n(\bar{X} - \mu)^2}{\sigma^2} + \frac{(n-1)S^2}{\sigma^2} - n \ln \left(\frac{(n-1)S^2}{n\sigma^2} \right) - n \leq d_n(1 - \alpha) \right\}$$

$$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2 \text{ proti } H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$$

Porovnanie testov a oblastí spoľahlivosti

Presné testy, presné oblasti spoľahlivosti

Upravený Mood

$$\mathcal{R} = \{(\mu, \sigma^2) : \sqrt{n}|\bar{X} - \mu|/\sigma \leq u(1 - \alpha_1/2) \\ \wedge (n-1)S^2/\sigma^2 \geq \chi_{n-1}^2(\alpha_2)\}$$

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Návrh upraviť LRT

$$\mathcal{R} = \left\{ (\mu, \sigma^2) : \frac{n(\bar{X} - \mu)^2}{\sigma^2} + \frac{(n-1)S^2}{\sigma^2} - n \ln \left(\frac{(n-1)S^2}{n\sigma^2} \right) - n \leq d_n(1 - \alpha) \right\}$$

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Porovnanie testov a oblastí spoľahlivosti

Presné testy, presné oblasti spoľahlivosti

Upravený Mood

$$\mathcal{R} = \{(\mu, \sigma^2) : \sqrt{n}|\bar{X} - \mu|/\sigma \leq u(1 - \alpha_1/2) \\ \wedge (n-1)S^2/\sigma^2 \geq \chi_{n-1}^2(\alpha_2)\}$$

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Návrh upraviť LRT

$$\mathcal{R} = \left\{ (\mu, \sigma^2) : \frac{n(\bar{X} - \mu)^2}{\sigma^2} + \frac{(n-1)S^2}{\sigma^2} - n \ln \left(\frac{(n-1)S^2}{n\sigma^2} \right) - n \leq d_n^*(1 - \alpha) \right\}$$

Presné testy, presné oblasti spoľahlivosti

Upravený Mood

$$\mathcal{R} = \{(\mu, \sigma^2) : \sqrt{n}|\bar{X} - \mu|/\sigma \leq u(1 - \alpha_1/2) \\ \wedge (n-1)S^2/\sigma^2 \geq \chi_{n-1}^2(\alpha_2)\}$$

$$- 1 - \alpha = (1 - \alpha_1)(1 - \alpha_2) \text{ a } \alpha_1 = \alpha_2 = 1 - \sqrt{1 - \alpha}$$

Návrh upraviť LRT

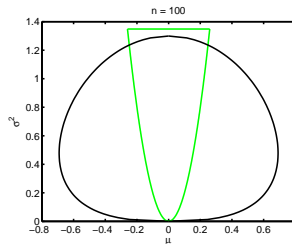
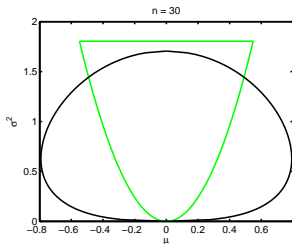
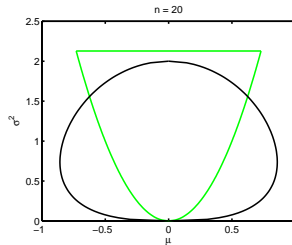
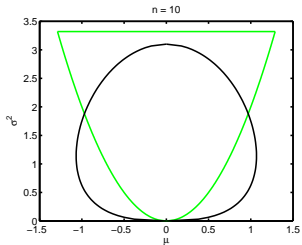
$$\mathcal{R} = \left\{ (\mu, \sigma^2) : \frac{n(\bar{X} - \mu)^2}{\sigma^2} - n \ln \left(\frac{(n-1)S^2}{\sigma^2} \right) \leq d_n^*(1 - \alpha) \right\}$$

$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2$ proti $H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$

Porovnanie testov a oblastí spoľahlivosti
Porovnanie testov a oblastí spoľahlivosti

Porovnanie: Silofunkcie testov, obsahy oblastí spoľahlivosti

Tvary oblastí spoľahlivosti pre $\alpha = 0,05$ a rôzne rozsahy výberu



$$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2 \text{ proti } H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$$

Porovnanie testov a oblastí spoľahlivosti
 Porovnanie testov a oblastí spoľahlivosti

Porovnávacia štúdia

Porovnávame:

- obsahy:

(pre $\alpha \in \{0,10; 0,05; 0,01\}$ a $n \in \{10, 20, 30, 100\}$)

obsah = [] $\times S^3$

- silofunkcie

(pre $\alpha = 0,05$ a $n \in \{10, 20\}$)

Silofunkcie testov závisia od $(\mu_1 - \mu_0)^2$, σ_1^2/σ_0^2

- uvažovali sme $\mu_0 = 0$ a $\sigma_0^2 = 1$
- meníme μ_1 pri pevnom $\sigma_1^2 = 1$
- meníme σ_1^2 pri pevnom $\mu_1 = 0$
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Porovnanie testov a oblastí spoľahlivosti
Porovnanie testov a oblastí spoľahlivosti

Obsahy $(1 - \alpha)$ -spoľahlivých oblastí

Metóda	mMood	mLRT	mMood	mLRT	mMood	mLRT
n	$\alpha = 0, 10$		$\alpha = 0, 05$		$\alpha = 0, 01$	
10	3,6148	3,6997	5,7058	5,2645	13,9660	10,8755
20	1,4842	2,1979	2,0706	2,7300	3,8678	4,2086
30	0,9884	1,8170	1,3192	2,1489	2,2436	2,9915
100	0,3774	1,3223	0,4669	1,4291	0,6793	1,6792
10	3,4181		5,3218		12,7547	
20	1,4753		2,0463		3,7814	
30	0,9884		1,3172		2,2275	
100	0,3689		0,4613		0,6774	

mMood*

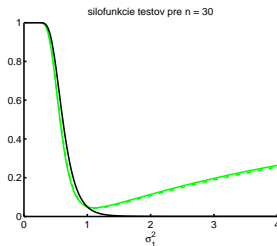
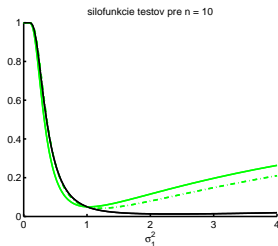
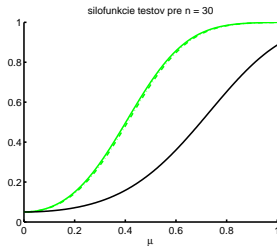
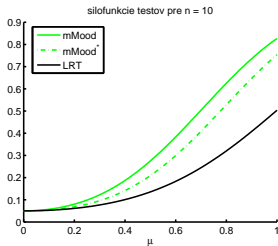
$$\mathcal{R} = \left\{ (\mu, \sigma^2) : \frac{\sqrt{n}|\bar{X} - \mu|}{\sigma} \leq u(1 - \alpha_1/2) \wedge \frac{(n-1)S^2}{\sigma^2} \geq \chi_{n-1}^2(\alpha_2) \right\}$$

$H_0 : \mu = \mu_0 \wedge \sigma^2 = \sigma_0^2$ proti $H_1 : \mu \neq \mu_0 \vee \sigma^2 < \sigma_0^2$

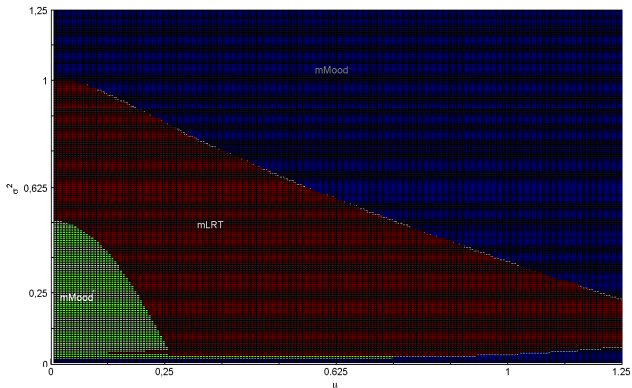
Porovnanie testov a oblastí spoľahlivosti
Porovnanie testov a oblastí spoľahlivosti

Plochy oblastí

Silofunkcie pre $\alpha = 0,05$



Silofunkcie pre rôzne μ_1, σ_1^2 pre $\alpha = 0,05; n = 10$



Ďakujem za pozornosť!